ON RECURSIVE ESTIMATION FOR TIME VARYING AUTOREGRESSIVE PROCESSES

ERIC MOULINES*, PIERRE PRIOURET†, AND FRANÇOIS ROUEFF*

*GET/Télécom Paris, CNRS LTCI 46 rue Barrault, 75634 Paris Cedex 13, France

[†]Laboratoire de Probabilités Université Paris VI 4, Place Jussieu, 75252 Paris cedex 05, France

ABSTRACT. This paper focuses on recursive estimation of time varying autoregressive processes in a non-parametric setting. The stability of the model is revisited and uniform results are provided when the time-varying autoregression parameters belong to appropriate smoothness classes. An adequate normalization for the correction term used in the recursive estimation procedure allows for very mild assumptions on the innovations distributions. The rate of convergence of the pointwise estimates are shown to be minimax in β -Lipschitz classes for $0 < \beta \le 1$. For $1 < \beta \le 2$, this property no longer holds. This can be seen by using an asymptotic expansion of the estimation error. A bias reduction method is then proposed for recovering the minimax rate.

1. Introduction

Suppose that we have real-valued observations $(X_{1,n}, X_{2,n}, \dots, X_{n,n})$ from a time-varying autoregressive model (TVAR)

(1)
$$X_{k,n} = \sum_{i=1}^{d} \theta_i((k-1)/n)X_{k-i,n} + \sigma(k/n)\epsilon_{k,n}, \qquad k = 1, \dots, n,$$

Date: February 24, 2005 (2nd revision).

AMS 2000 MSC. Primary 62M10, 62G08, 60J27. Secondary 62G20.

 $\label{eq:keywords:locally stationary processes, non-parametric estimation, recursive estimation, time-varying autoregressive model.$

where

- $\{\epsilon_{k,n}\}_{1\leq k\leq n}$ is a triangular array of real valued random variables referred to as the (normalized) *innovations*,
- $\boldsymbol{\theta}(t) := [\theta_1(t) \dots \theta_d(t)]^T$, $t \in [0, 1]$, is a d-dimensional vector referred to as the *local autoregression vector*,
- $\sigma(t)$, $t \in [0,1]$ is a non-negative number referred to as the *local* innovation standard deviation.

This recurrence equation may be more compactly written as

(2)
$$X_{k,n} = \boldsymbol{\theta}_{k-1,n}^T \mathbf{X}_{k-1,n} + \sigma_{k,n} \epsilon_{k,n}, \qquad k = 1, \dots, n,$$

where

$$\mathbf{X}_{k,n} := [X_{k,n} \ X_{k-1,n} \ \dots \ X_{k-d+1,n}]^T,$$

$$\boldsymbol{\theta}_{k,n} := \boldsymbol{\theta}(k/n) = [\theta_1(k/n) \ \theta_2(k/n) \ \dots \ \theta_d(k/n)]^T \quad \text{and} \quad \sigma_{k,n} = \sigma(k/n).$$

TVAR models have been used for modeling data whose spectral content varies along time (see for instance Subba Rao (1970), Hallin (1978) and Grenier (1983) for early references). TVAR models are also closely related to the general class of locally stationary processes (see Dahlhaus (1996a), Dahlhaus (1996b) and Dahlhaus (1997) and Remark 3 below for definitions and properties).

In this paper, we focus on the estimation of the functions $t \mapsto \boldsymbol{\theta}(t)$ (we leave aside $\sigma(t)$ for brevity) from the observations $\{\mathbf{X}_{0,n}, X_{k,n}, k \geq 1\}$ (here we add the initial conditions $\mathbf{X}_{0,n}$ in the observations set for convenience). This problem is reminiscent of non-parametric curve estimation on a fixed design, a problem which has received a considerable attention in the literature. A natural approach consists in using a stationary method on short overlapping segments of the time series (see for instance Dahlhaus and Giraitis (1998)). An alternative approach, first investigated by Belitser (2000) for first order TVAR models consists in estimating the regression function recursively in time. More precisely, at a given time $t \in (0,1)$, only observations that have been observed before time t are used in the definition of the estimator: $\hat{\boldsymbol{\theta}}_n(t) = \hat{\boldsymbol{\theta}}_n(t, \mathbf{X}_{0,n}, X_{1,n}, \dots, X_{[nt],n})$, where [x] denotes the integer part of x. This approach is useful when the observations must be processed on line (see e.g. Ljung and Soderström (1983), Solo and Kong

(1995), Kushner and Yin (1997)). We focus in this contribution on the Normalized Least Square algorithm (NLMS) which is a specific example of a recursive identification algorithm, defined as follows.

$$\widehat{\boldsymbol{\theta}}_{0,n}(\mu) := 0,$$

$$(3) \qquad \widehat{\boldsymbol{\theta}}_{k+1,n}(\mu) := \widehat{\boldsymbol{\theta}}_{k,n}(\mu) + \mu \left(X_{k+1,n} - \widehat{\boldsymbol{\theta}}_{k,n}^T(\mu) \, \mathbf{X}_{k,n} \right) \frac{\mathbf{X}_{k,n}}{1 + \mu |\mathbf{X}_{k,n}|^2},$$

where k goes from zero to n-1, μ is referred to as the step-size and $|\cdot|$ denotes the Euclidean norm. At each iteration of the algorithm, the parameter estimates is updated by moving in the direction of the gradient of the instantaneous estimate $(X_{k+1,n} - \boldsymbol{\theta}^T \mathbf{X}_{k,n})^2$ of the local mean square error $\mathbb{E}[(X_{k+1,n} - \boldsymbol{\theta}^T \mathbf{X}_{k,n})^2]$. The normalization $(1 + \mu |\mathbf{X}_{k,n}|^2)^{-1}$ is a safeguard again large values of the norm of the regression vector and allows for very mild assumptions on the innovations (see (A1) below) compared with the LMS, which typically requires much stronger assumptions (see Priouret and Veretennikov (1995)). Extensions to more sophisticated iterative rule, e.g. the so called recursive least-square (RLS) algorithm, are currently under investigation. We define a pointwise estimate of $t \mapsto \boldsymbol{\theta}(t)$ as a simple interpolation of $\hat{\theta}_{k,n}(\mu)$, $k = 1, \ldots, n$, i.e.

(4)
$$\widehat{\boldsymbol{\theta}}_n(t;\mu) := \widehat{\boldsymbol{\theta}}_{[tn],n}(\mu), \quad t \in [0,1], \ n \ge 1.$$

Observe that, for all $t \in [0,1]$, $\widehat{\boldsymbol{\theta}}_n(t;\mu)$ is a function of $\mathbf{X}_{0,n}$, $\{X_{l,n}, l = 1, \ldots, [tn]\}$ and μ .

The paper is organized as follows. In section 2, we introduce the assumptions and state the main results of this paper, namely, uniform risk bounds for $\hat{\theta}_n$, a lower bound of the minimax L^2 -risk and precise approximations of the risk for $\hat{\theta}_n$. We also discuss a technique to correct the bias of the estimator. In Section 3 the basic results used along the paper for dealing with the weak dependence are presented. In Section 4, a four steps proof of the uniform upper bound for the L^p risk of $\hat{\theta}_n$ is given. Section 5 then provides the proof of the minimax lower bound. In Section 6, further approximation results are given from which we obtain the proofs of the risk approximations for $\hat{\theta}_n$ stated in Section 2.

2. Main results

The space of $m \times n$ matrices is embedded with the operator norm associated to the Euclidean norm, which we denote by

$$(5) |A| := \sup_{x \in \mathbb{R}^n, |x|=1} |Ax|.$$

Observe that for a row or column vector, its Euclidean norm coincides with its operator norm. For any random variable **Z** in a normed space $(\mathbf{Z}, |\cdot|)$, we denote by $\|\mathbf{Z}\|_p := (\mathbb{E}|\mathbf{Z}|^p)^{\frac{1}{p}}$ its L^p -norm. Throughout the paper, it is assumed that,

(A1) for all $n \geq 1$, the random variables $\{\epsilon_{k,n}\}_{1 \leq k \leq n}$ are independent, have zero mean and unit variance and are independent of the initial conditions $\mathbf{X}_{0,n}$. In addition, $\sup_{n \geq 1} \|\mathbf{X}_{0,n}\|_q < \infty$ and $\epsilon_q^* := \sup_{1 \leq k \leq n} \|\epsilon_{k,n}\|_q < \infty$,

where the moment order $q \geq 2$ will be set depending on the context. The triangular array of random variables $\{X_{k,n}, 1 \leq k \leq n\}$ defined by (1) is parameterized by $(\boldsymbol{\theta}, \sigma)$. To keep track of the dependence in $(\boldsymbol{\theta}, \sigma)$, for all random variable Z defined as a function of these variables, we shall adopt the notation convention $\mathbb{E}_{\boldsymbol{\theta},\sigma}[Z]$ for the expectation of Z. In the case of a random element \mathbf{Z} taking its values in the normed space $(\mathbf{Z}, |\cdot|)$, its L^p -norm will be denoted by

$$\|\mathbf{Z}\|_{p,\boldsymbol{\theta},\sigma} := (\mathbb{E}_{\boldsymbol{\theta},\sigma}|\mathbf{Z}|^p)^{1/p}$$
.

A classical problem in non-parametric estimation is to derive uniform bounds for the pointwise L^p risk $\|\widehat{\boldsymbol{\theta}}_n(t;\mu) - \boldsymbol{\theta}(t)\|_{p,\boldsymbol{\theta},\sigma}$ for $(\boldsymbol{\theta},\sigma)$ in some appropriate classes of functions. In the sequel we denote by

(6)
$$\theta(z;t) := 1 - \sum_{j=1}^{d} \boldsymbol{\theta}_{j}(t) z^{j}, \quad z \in \mathbb{C},$$

the local time-varying autoregressive polynomial associated to $\boldsymbol{\theta}$ at point t. The function classes that will be considered rely on two kinds of properties. First, the roots of the time-varying autoregressive polynomial associated to $\boldsymbol{\theta}$ are required to stay away from the unit disk. Second, as in function estimation from noisy data, $\boldsymbol{\theta}$ and $\boldsymbol{\sigma}$ are supposed to be smooth in some appropriate sense. The first condition is unusual in the non-parametric function estimation setting and deserves some elaboration. Let us recall

some elementary facts from the theory of autoregressive processes. The process $\{Z_k\}$ is an AR(d) process with parameters $\vartheta \in \mathbb{R}^d$ and $\varsigma^2 > 0$ if $\{Z_k\}$ is second order stationary and satisfies the following difference equation

(7)
$$Z_k = \sum_{j=1}^d \vartheta_j Z_{k-j} + \varepsilon_k, \quad k \in \mathbb{Z},$$

where $\{\varepsilon_k\}_{k\in\mathbb{Z}}$ is a centered white noise with variance ς^2 . A sufficient and necessary condition for the existence of $\{Z_k\}$ is that the autoregressive polynomial $z\mapsto \vartheta(z):=1-\sum_{j=1}^d \vartheta_j\,z^j$ does not vanish on the unit circle (see Brockwell and Davis (1991)). In this case, the stationary solution $\{Z_k\}$ to (7) is unique and there exists a sequence $\{\psi_k\}$, such that $\sum_l |\psi_l| < \infty$ and $Z_k = \sum_l \psi_l \epsilon_{k-l}$ for all $k \in \mathbb{Z}$. Furthermore the sequence $\{\psi_l\}$ is causal, *i.e.* $\psi_l = 0$ for all l < 0, if and only if the function $z \mapsto \vartheta(z)$ does not vanish on the disk $|z| < \rho^{-1}$, for some $\rho < 1$. This motivates the following definitions in the time varying setting. For $\rho > 0$, we denote

(8)
$$S(\rho) := \left\{ \boldsymbol{\theta} : [0,1] \to \mathbb{R}^d, \ \theta(z;t) \neq 0 \text{ for all } |z| < \rho^{-1} \text{ and } t \in [0,1] \right\}.$$

Concerning the smoothness condition, different classes of functions can be considered. In the original paper by (Dahlhaus, 1996b), it is assumed that the functions $t \mapsto \boldsymbol{\theta}(t)$ and $t \mapsto \sigma(t)$ are Lipschitzian. In this paper, we consider a wider range of smoothness classes which are now introduced. For any $\beta \in (0,1]$, denote the β -Lipschitz semi-norm of a mapping $\mathbf{f}:[0,1]\mapsto \mathbb{R}^l$ by

$$|\mathbf{f}|_{\Lambda,\beta} = \sup_{t \neq s} \frac{|\mathbf{f}(t) - \mathbf{f}(s)|}{|t - s|^{\beta}}.$$

and define for $0 < L < \infty$, the β -Lipschitz ball

(9)
$$\Lambda_l(\beta, L) := \left\{ \mathbf{f} : [0, 1] \to \mathbb{R}^l, |\mathbf{f}|_{\Lambda, \beta} \le L, \sup_{t \in [0, 1]} |\mathbf{f}(t)| \le L \right\}.$$

For all $\beta > 1$ the β -Lipschitz balls are classically generalized as follows. Let $k \in \mathbb{N}$ and $\alpha \in (0,1]$ be uniquely defined by $\beta = k + \alpha$. Then we define

(10)
$$\Lambda_l(\beta, L) := \left\{ \mathbf{f} : [0, 1] \to \mathbb{R}^l, |\mathbf{f}^{(k)}|_{\Lambda, \alpha} \le L, \sup_{t \in [0, 1]} |\mathbf{f}(t)| \le L \right\}.$$

where $\mathbf{f}^{(k)}$ is the derivative of order k.

For all $\beta > 0$, L > 0, $0 < \rho < 1$, and $0 < \sigma_{-} \le \sigma_{+} < \infty$, we define

$$\mathcal{C}(\beta, L, \rho, \sigma_{-}, \sigma_{+}) := \{ (\boldsymbol{\theta}, \sigma) : \boldsymbol{\theta} \in \Lambda_{d}(\beta, L) \cap \mathcal{S}(\rho), \sigma : [0, 1] \to [\sigma_{-}, \sigma_{+}] \}.$$

We will simply write \mathcal{C} whenever no confusion is possible. It is interesting to observe that, for particular choices of L and ρ , \mathcal{C} reduces to the more conventional smoothness class $\{(\boldsymbol{\theta}, \sigma) : \boldsymbol{\theta} \in \Lambda_d(\beta, L), \sigma : [0, 1] \to [\sigma_-, \sigma_+]\}$. This follows from the following lemma.

Lemma 1. For all positive ρ , we have

(11)
$$B(1/\sqrt{\rho^{-2} + \dots + \rho^{-2d}}) \subseteq \mathcal{S}(\rho) \subseteq B((1+\rho)^d - 1),$$

where B(a) is the sup-norm ball $\{\boldsymbol{\theta}: [0,1] \to \mathbb{R}^d, \sup_{t \in [0,1]} |\boldsymbol{\theta}(t)| \le a \}$.

Proof. Note that $\theta(t;0) = 1$ for all $t \in [0,1]$. Let $\lambda_1(t), \ldots, \lambda_d(t)$ be the reciprocals of the roots of the polynomial $z \mapsto \theta(z;t)$. Hence $\theta \in \mathcal{S}(\rho)$ iff, for all $t \in [0,1]$ and $k = 1, \ldots, d$, $|\lambda_k(t)| \leq \rho$. For all $t \in [0,1]$ and $k = 1, \ldots, d$, using the Cauchy-Schwarz inequality, we have

$$1 = \left| \sum_{i=1}^d \theta_i(t) \lambda_k^{-i}(t) \right| \le |\boldsymbol{\theta}(t)| \left(\sum_{i=1}^d |\lambda_k(t)|^{-2i} \right)^{1/2}.$$

Hence the first inclusion in (11). We further have

$$\theta(z;t) = \prod_{k=1}^{d} (1 - \lambda_k(t) z).$$

The coefficients of θ are then given by

(12)
$$\theta_k(t) = (-1)^k \sum_{1 \le i_1 < \dots < i_k \le d} \lambda_{i_1}(t) \dots \lambda_{i_k}(t), \quad k = 1, \dots, d.$$

A simple computation then gives the second inclusion in (11).

Remark 1. In the case d=1 considered in Belitser (2000), $S(\rho)=B(\rho)$.

We are now in a position where we can state the main results of this paper. We first provide a uniform upper bound of the pointwise L^p risk of the NLMS estimator.

Theorem 2. Assume (A1) with $q \ge 4$ and let $p \in [1, q/3)$. Let $\beta \in (0, 1]$, L > 0, $0 < \rho < 1$, and $0 < \sigma_- \le \sigma_+$. Then, there exists $M, \delta > 0$ and $\mu_0 > 0$ such that, for all $\mu \in (0, \mu_0]$, $n \ge 1$, $t \in (0, 1]$ and $(\theta, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_-, \sigma_+)$,

(13)
$$\|\widehat{\boldsymbol{\theta}}_n(t;\mu) - \boldsymbol{\theta}(t)\|_{p,\boldsymbol{\theta},\sigma} \le M \left(|\boldsymbol{\theta}(0)| (1 - \delta \mu)^{tn} + \sqrt{\mu} + (n\mu)^{-\beta} \right).$$

Corollary 3. For all $\eta \in (0,1)$ and $\alpha > 0$, there exists M > 0 such that, for all $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_-, \sigma_+)$,

$$\sup_{t \in [\eta,1]} \left\| \widehat{\boldsymbol{\theta}}_n \left[t; \alpha n^{-2\beta/(1+2\beta)} \right] - \boldsymbol{\theta}(t) \right\|_{p,\boldsymbol{\theta},\sigma} \leq M \, n^{-\frac{\beta}{1+2\beta}}.$$

The upper bound in (13) has three terms. Anticipating on what will be said in the proof section (Section 4), the first term $|\boldsymbol{\theta}(0)| (1-\delta \mu)^{tn}$ reflects the forgetting of the initial error of the NLMS estimator. The second term is the so-called lag-noise term, which accounts for the fluctuation of the recursive estimator which would occur even if $t \mapsto \boldsymbol{\theta}(t)$ is constant. The third term controls the error involved by time evolution of $\boldsymbol{\theta}(t)$ and mainly relies on the smoothness exponent β . Corollary 3 is then obtained by choosing the step-size in order to minimize this upper bound. Observing that the first term is negligible for $t \geq \eta > 0$ and balancing the two remaining ones yields $\mu \propto n^{\frac{2\beta}{1+2\beta}}$. This corollary says that, for all $\beta \in (0,1]$, under β -Lipschitz assumption, for $t \in (0,1]$, the L^p risk of the NLMS estimator at point t has rate $n^{-\frac{\beta}{1+2\beta}}$.

We now provide a lower bound of the L^2 -risk for any estimator $\hat{\boldsymbol{\delta}}_n$ of $\boldsymbol{\theta}(t)$ computed from observations $\mathbf{X}_{0,n}, X_{1,n}, \ldots, X_{n,n}$. Let us stress that this lower bound is not restricted to recursive estimators, *i.e.* we do not require $\hat{\boldsymbol{\delta}}_n$ to depend only on $\mathbf{X}_{0,n}, X_{1,n}, \ldots, X_{[nt],n}$. Denote by

$$\mathrm{MSEM}_{\boldsymbol{\theta},\sigma}(\widehat{\boldsymbol{\delta}}_n,t) := \mathbb{E}_{\boldsymbol{\theta},\sigma}\left[(\widehat{\delta}_n - \boldsymbol{\theta}(t))(\widehat{\delta}_n - \boldsymbol{\theta}(t))^T\right] \ .$$

the Mean Square Error Matrix (MSEM) at $t \in [0, 1]$. Consider the following assumption.

(A2) For all $n \in \mathbb{N}$ and $1 \le k \le n$, $\epsilon_{k,n}$ has an absolutely continuous density $p_{k,n}$ w.r.t. the Lebesgue measure whose derivative $\dot{p}_{k,n}$ satisfies

$$\mathbb{E}\left[\frac{\dot{p}_{k,n}}{p_{k,n}}(\epsilon_{k,n})\right] = 0 \quad \text{and} \quad \mathcal{I}_{\epsilon} := \sup_{1 \leq k \leq n} \mathbb{E}\left[\left(\frac{\dot{p}_{k,n}}{p_{k,n}}(\epsilon_{k,n})\right)^2\right] < \infty.$$

We have the following result, whose proof is postponed to Section 5.

Theorem 4. Assume (A1) with q = 2 and (A2). Let $\beta > 0$, L > 0, $\rho \in (0,1)$, $0 < \sigma_{-} \leq \sigma_{+}$. Then there exists $\alpha > 0$ such that, for all $n \geq 1$, $t \in [0,1]$, and for all estimator $\hat{\delta}_{n} := \hat{\delta}_{n}(\mathbf{X}_{0,n}, X_{1,n}, \dots, X_{n,n}) \in \mathbb{R}^{d}$,

(14)
$$\inf_{\|\mathbf{u}\|=1} \sup_{(\boldsymbol{\theta},\sigma)\in\mathcal{C}} \mathbf{u}^T \mathbf{MSEM}_{\boldsymbol{\theta},\sigma}(\widehat{\boldsymbol{\delta}}_n,t) \mathbf{u} \ge \alpha n^{\frac{-2\beta}{1+2\beta}}.$$

where $\mathcal{C} := \mathcal{C}(\beta, L, \rho, \sigma_-, \sigma_+)$.

Remark 2. Note that $\mathbf{u}^T \mathrm{MSEM}_{\boldsymbol{\theta},\sigma}(\widehat{\boldsymbol{\delta}}_n,t)\mathbf{u}$ is the mean square error of $\mathbf{u}^T \widehat{\boldsymbol{\delta}}_n$ for estimating $\mathbf{u}^T \boldsymbol{\theta}(t)$.

Corollary 3 and Theorem 4 show that, under (A1) with q > 6 and (A2), the L^p error rate is minimax for $p \geq 2$ within the class $\mathcal{C}(\beta, L, \rho, \sigma_-, \sigma_+)$ if $\beta \in (0, 1]$. The question arises whether the upper bound derived in Theorem 2 generalizes for $\beta > 1$ in such a way that, as in Corollary 3, for an appropriate step-size $\mu(n)$, $\widehat{\boldsymbol{\theta}}_n(t;\mu(n))$ achieves the rate of the lower bound derived in Theorem 4. It turns out that this is not the case, except in a very particular situation, namely when $\boldsymbol{\theta}$ is the constant function. This may be shown by using precise approximations of the risk, completing the upper bound given in Theorem 2. Such approximations primarily rely on the fact that, as n tend to infinity, the local sample $\mathbf{X}_{k,n}$ of the TVAR process approximately has the same second-order statistics as the stationary AR(d) process with parameter $(\boldsymbol{\theta}(k/n), \sigma(k/n))$. In the following we provide a precise statement of this fact and then state the approximations of the risk. For this purpose, we need to introduce further notations. For $\beta > 0$, L > 0, $\rho \in (0,1)$ and $0 < \sigma_- \le \sigma_+$ we let

(15)
$$\mathcal{C}^{\star}(\beta, L, \rho, \sigma_{-}, \sigma_{+}) := \{ (\boldsymbol{\theta}, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_{-}, \sigma_{+}) : \sigma \in \Lambda_{1}(\beta, \sigma_{+}) \}.$$

We use the shorthand notation C^* when no confusion is possible. The obvious relation

$$\mathcal{C}(\beta, L, \rho, \sigma_{-}, \sigma_{+}) \supseteq \mathcal{C}^{\star}(\beta, L, \rho, \sigma_{-}, \sigma_{+}) \supseteq \mathcal{C}(\beta, L, \rho, \sigma_{+}, \sigma_{+})$$

implies that Theorem 2 and Theorem 4 are still valid when replacing C by C^* . Following the formula of the spectral densities of stationary AR(d)

processes, we respectively denote

(16)
$$f(\lambda; t, \boldsymbol{\theta}, \sigma) := \frac{\sigma^2(t)}{2\pi} \left| \theta(e^{i\lambda}; t) \right|^{-2}, \quad -\pi \le \lambda \le \pi,$$

(17)
$$[\Sigma(t,\boldsymbol{\theta},\sigma)]_{k,l} := \int_{-\pi}^{\pi} e^{i\lambda(k-l)} f(\lambda;t,\boldsymbol{\theta},\sigma) d\lambda, \quad 1 \leq k,l \leq d,$$

the local spectral density function and the local d-dimensional covariance matrix associated to $(\boldsymbol{\theta}, \sigma)$ at point t. The covariance matrix $\mathbb{E}_{\boldsymbol{\theta}, \sigma}[\mathbf{X}_{k,n}\mathbf{X}_{k,n}^T]$ can be approximated by the local covariance matrix at point k/n as follows.

Proposition 5. Assume (A1) with $q \geq 2$. Let $\beta \in (0,1]$, L > 0, $0 < \rho < \tau < 1$, and $0 < \sigma_{-} \leq \sigma_{+}$. Then there exists M > 0 such that, for all $1 \leq k \leq n$ and $(\theta, \sigma) \in C^{\star}(\beta, L, \rho, \sigma_{-}, \sigma_{+})$,

$$\left| \mathbb{E}_{\boldsymbol{\theta}, \sigma} [\mathbf{X}_{k,n} \mathbf{X}_{k,n}^T] - \Sigma(k/n, \boldsymbol{\theta}, \sigma) \right| \leq M \left(\tau^k \left| \mathbb{E} [\mathbf{X}_{0,n} \mathbf{X}_{0,n}^T] - \Sigma(0, \boldsymbol{\theta}, \sigma) \right| + n^{-\beta} \right).$$

Remark 3. This approximation result can serve as an illustration of how the TVAR process fits in the locally stationary setting introduced Dahlhaus (1996a). Observe that if $\beta = 1$ and $\mathbb{E}[\mathbf{X}_{0,n}\mathbf{X}_{0,n}^T] = \Sigma(0, \boldsymbol{\theta}, \sigma)$, the rate for the approximation error between the local sample covariance matrix $\mathbb{E}_{\boldsymbol{\theta},\sigma}[\mathbf{X}_{k,n}\mathbf{X}_{k,n}^T]$ and its local stationary approximation $\Sigma(k/n,\boldsymbol{\theta},\sigma)$ is n^{-1} , which coincides with the approximation rate required in the locally stationary setting introduced in Dahlhaus (1996a). Precise conditions upon which a TVAR process is locally stationary are given in Dahlhaus (1996b).

We obtain the following computational approximation of the pointwise MSEM for $\widehat{\boldsymbol{\theta}}_n$.

Theorem 6. Assume (A1) with q > 11. Let $\beta \in (0,1]$, L > 0, $\rho < 1$, and $0 < \sigma_{-} \leq \sigma_{+}$ and let $(\theta, \sigma) \in \mathcal{C}^{*}(\beta, L, \rho, \sigma_{-}, \sigma_{+})$. Let $t \in (0,1]$ and assume that there exists $\theta_{t,\beta} \in \mathbb{R}^{d}$, L' > 0 and $\beta' > \beta$ such that, for all $u \in [0,t]$,

(18)
$$|\boldsymbol{\theta}(u) - \boldsymbol{\theta}(t) - \boldsymbol{\theta}_{t,\beta} (t - u)^{\beta}| \le L' (t - u)^{\beta'}.$$

Then there exists M>0 and $\mu_0>0$ such that for all $\mu\in(0,\mu_0]$ and $n\geq 1,$

(19)
$$\left| \operatorname{MSEM}_{\boldsymbol{\theta},\sigma} \left(\widehat{\boldsymbol{\theta}}_{n}(t;\mu) - \frac{\Gamma(\beta+1)}{(\mu\,n)^{\beta}} \Sigma^{-\beta}(t,\boldsymbol{\theta},\sigma) \boldsymbol{\theta}_{t,\beta} \right) - \mu \, \frac{\sigma^{2}(t)}{2} \, I \right|^{1/2} \leq M \left(\sqrt{\mu} \left(\sqrt{\mu} + (\mu\,n)^{-\beta/2} \right) + (\mu\,n)^{-\beta} \left((\mu\,n)^{-\beta} + (\mu\,n)^{\beta-\beta'} + \sqrt{\mu} \right) \right),$$

where Γ is the usual Gamma function and I is the identity matrix with size $d \times d$

Remark 4. Let $\alpha \in \mathbb{R}$. The α -fractional power of a diagonal matrix D with positive diagonal entries is the diagonal matrix D^{α} obtained by raising the diagonal entries to the power α . The α -fractional power of a symmetric positive-definite matrix $A = UDU^T$, where U is unitary and D is diagonal, is then defined by $A^{\alpha} = UD^{\alpha}U^T$.

Using (19), as $(\mu + (\mu n)^{-1}) \to 0$, we have the following asymptotic approximation of the MSEM.

$$MSEM_{\boldsymbol{\theta},\sigma}\left(\widehat{\boldsymbol{\theta}}_n(t;\mu) - \Gamma(\beta+1) (\mu n)^{-\beta} \Sigma^{-\beta}(t,\boldsymbol{\theta},\sigma) \boldsymbol{\theta}_{t,\beta}\right) = \mu \frac{\sigma^2(t)}{2} I(1+o(1)).$$

If $\beta \leq 1$, this gives the leading term of an asymptotic expansion of the MSEM, which allows to compare the performance of $\widehat{\boldsymbol{\theta}}_n$ with other estimators achieving the minimax rate. The deterministic correction $\Gamma(\beta+1)(\mu n)^{-\beta}\Sigma^{-\beta}(t,\boldsymbol{\theta},\sigma)\boldsymbol{\theta}_{t,\beta}$ can be interpreted as the main term of the bias and the term $\mu \frac{\sigma^2(t)}{2}I$ as the main term of the covariance matrix.

Remark 5. An essential ingredient for proving Theorem 6 is to approximate the expectation of the term $\mathbf{X}_{k,n}^T \mathbf{X}_{k,n}/(1+\mu|\mathbf{X}_{k,n}|^2)$ appearing in (3). Roughly speaking, for small enough μ , a good approximation is $\mathbb{E}_{\boldsymbol{\theta},\sigma}[\mathbf{X}_{k,n}^T \mathbf{X}_{k,n}]$ which itself is well approximated by $\Sigma(k/n,\boldsymbol{\theta},\sigma)$ by using Proposition 5. If one replaces the normalization factor $1+\mu|\mathbf{X}_{k,n}|^2$ by the more classical (in the stochastic tracking literature) $1+|\mathbf{X}_{k,n}|^2$, the computation of the deterministic approximation would be much more involved as the normalization does not reduce to one as μ tend to zero.

Remark 6. Equation (18) is for instance valid if $\boldsymbol{\theta}$ behaves as a sum of non-entire power laws at the left of point t, say $\boldsymbol{\theta}(u) = \boldsymbol{\theta}_{t,1/2} \sqrt{t-u} + O(t-u)$. Another simple case consists in assuming that $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}^{\star}(\beta', L, \rho, \sigma_{-}, \sigma_{+})$ for $\beta' > 1$. Then (18) is obtained with $\beta = 1$ and $\boldsymbol{\theta}_{t,\beta} = -\dot{\boldsymbol{\theta}}(t)$ by using a first order Taylor expansion. Hence, in this case, the main terms of the MSEM are of order $\mu + (\mu n)^{-2}$ unless $\dot{\boldsymbol{\theta}}(t) = 0$ in which case the deterministic correction in the MSEM vanishes. This implies that the estimator $\hat{\boldsymbol{\theta}}_n(t;\mu)$ cannot achieve the minimax rate obtained in Theorem 4 in the class $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}^{\star}(\beta', L, \rho, \sigma_{-}, \sigma_{+})$ for $\beta' > 1$ unless $\boldsymbol{\theta}$ is a constant function.

If the smoothness exponent belongs to (1,2], the following result applies.

Theorem 7. Assume (A1) with q > 4 and let $p \in [1, q/4)$. Let $\beta \in (1, 2]$, L > 0, $\rho \in (0, 1)$, and $0 < \sigma_{-} \leq \sigma_{+}$. Then, for all $\eta \in (0, 1)$, there exists M > 0 and $\mu_{0} > 0$ such that, for all $(\theta, \sigma) \in C^{*}(\beta, L, \rho, \sigma_{-}, \sigma_{+})$, $t \in [\eta, 1]$, $n \geq 1$ and $\mu \in (0, \mu_{0}]$,

$$\left\|\widehat{\boldsymbol{\theta}}_n(t;\mu) - \boldsymbol{\theta}(t) + (\mu \, n)^{-1} \Sigma^{-1}(t,\boldsymbol{\theta},\sigma) \, \dot{\boldsymbol{\theta}}(t) \right\|_{p,\boldsymbol{\theta},\sigma} \leq M \left(\sqrt{\mu} + (\mu \, n)^{-\beta} + (\mu \, n)^{-2}\right).$$

Applying a technique inspired by the so-called Romberg method in numerical analysis (see e.g. Baranger (1991)), we are now able to propose a recursive estimator which achieves the minimax rates for $\beta \in (1,2]$. This estimator is obtained by combining the recursive estimators $\hat{\theta}(t;\cdot)$ associated to two different step-sizes. More precisely, let

$$\widetilde{\boldsymbol{\theta}}_n(t;\mu,\gamma) := \frac{1}{1-\gamma} \left(\widehat{\boldsymbol{\theta}}_n(t;\mu) - \gamma \widehat{\boldsymbol{\theta}}_n(t;\gamma \mu) \right).$$

where $\gamma \in (0,1)$. We obtain the following result.

Theorem 8. Assume (A1) with q > 4 and let $p \in [1, q/4)$. Let $\beta \in (1, 2]$, L > 0, $\rho < 1$, and $0 < \sigma_- \le \sigma_+$. For all $\eta \in (0, 1)$, there exists M > 0 and $\mu_0 > 0$ such that, for all $\gamma \in (0, 1)$, $(\theta, \sigma) \in C^*(\beta, L, \rho, \sigma_-, \sigma_+)$, $n \ge 1$ and $\mu \in (0, \mu_0]$,

(20)
$$\sup_{t \in [\eta, 1]} \left\| \tilde{\boldsymbol{\theta}}_n(t; \mu, \gamma) - \boldsymbol{\theta}(t) \right\|_{p, \boldsymbol{\theta}, \sigma} \le M \frac{1 + \gamma}{\gamma (1 - \gamma)} \left(\sqrt{\mu} + (\mu \, n)^{-\beta} + (\mu \, n)^{-2} \right).$$

Proof. Let $\eta \in (0,1)$, $\gamma \in (0,1)$ and $t \in [\eta,1]$. One easily checks that

$$\widetilde{\boldsymbol{\theta}}_n(t;\mu,\gamma) - \boldsymbol{\theta}(t) = (1-\gamma)^{-1} \Big(\widehat{\boldsymbol{\theta}}_n(t;\mu) - \boldsymbol{\theta}(t) + (\mu n)^{-1} \Sigma^{-1}(t,\boldsymbol{\theta},\sigma) \dot{\boldsymbol{\theta}}(t) - \gamma \Big(\widehat{\boldsymbol{\theta}}_n(t;\gamma\mu) - \boldsymbol{\theta}(t) + (\gamma\mu n)^{-1} \Sigma^{-1}(t,\boldsymbol{\theta},\sigma) \dot{\boldsymbol{\theta}}(t) \Big) \Big).$$

The Minkowski Inequality, Theorem 7 and the bounds $\sqrt{\gamma} < \gamma^{-\beta} \le \gamma^{-2}$ yield (20).

Corollary 9. For all $\gamma, \eta \in (0,1)$, $\beta \in (0,2]$ and $\alpha > 0$, there exists M > 0 such that, for all $(\theta, \sigma) \in C^*(\beta, L, \rho, \sigma_-, \sigma_+)$,

$$\sup_{t \in [\eta,1]} \left\| \tilde{\boldsymbol{\theta}}_n(t;\alpha \, n^{-2\beta/(1+2\beta)},\gamma) - \boldsymbol{\theta}(t) \right\|_{p,\boldsymbol{\theta},\sigma} \leq M \, \frac{1+\gamma}{\gamma(1-\gamma)} \, n^{-\frac{\beta}{1+2\beta}}.$$

3. Exponential stability of inhomogeneous difference equations

Let us consider a sequence $\{\mathbf{Z}_k, k \geq 0\}$ of random vectors satisfying the following inhomogeneous difference equation

(21)
$$\mathbf{Z}_k = A_k \mathbf{Z}_{k-1} + B_k \mathbf{U}_k, \quad k \ge 1,$$

where $\{\mathbf{U}_k, k \geq 1\}$ is a sequence of independent random vectors and $A = \{A_k, k \geq 1\}$, $B = \{B_k, k \geq 1\}$ are two sequences of deterministic matrices with appropriate dimensions. The pair (A, B) is said to be *exponentially* stable if there exist constants C > 0 and $\rho \in (0, 1)$ such that

(22)
$$\sup_{k\geq 0} \left| \prod_{l=k+1}^{k+m} A_l \right| \leq C\rho^m \quad \text{for all } m>0 \quad \text{and} \quad B^* := \sup_{k\geq 1} |B_k| < \infty,$$

with the convention that $\prod_{l=k}^{k+m} A_l := A_{k+m} A_{k+m-1} \dots A_k$. We clearly have

Proposition 10. Let $p \in [1, \infty]$. Suppose that (A, B) is exponentially stable. Then there exists a positive constant M only depending on C, ρ and B^* such that

$$\|\mathbf{Z}_k\|_p \le M \left(\mathbf{U}_p^* + \|\mathbf{Z}_0\|_p\right), \quad k \in \mathbb{N},$$

where $\mathbf{U}_p^{\star} := \sup_{k \geq 1} \|\mathbf{U}_k\|_p$.

Exponential stability implies exponential forgetting of the initial condition in the following setting. Let $(\mathsf{E},|\cdot|_\mathsf{E})$ and $(\mathsf{F},|\cdot|_\mathsf{F})$ be two normed spaces, let m be a positive integer and p a non-negative real number. We denote by $\mathrm{Li}(\mathsf{E},m,\mathsf{F};p)$ the linear space of mappings $\phi:\mathsf{E}^m\to\mathsf{F}$ for which there exists $\lambda_1\geq 0$ such that, for all (x_1,\ldots,x_m) and $(y_1,\ldots,y_m)\in\mathsf{E}^m$,

$$(23) \quad |\phi(x_1,\ldots,x_m) - \phi(y_1,\ldots,y_m)|_{\mathsf{F}} \le$$

$$\lambda_1 (|x_1 - y_1|_{\mathsf{E}} + \dots + |x_m - y_m|_{\mathsf{E}}) (1 + |x_1|_{\mathsf{E}}^p + |y_1|_{\mathsf{E}}^p + \dots + |x_m|_{\mathsf{E}}^p + |y_m|_{\mathsf{E}}^p).$$

This implies that there exists $\lambda_2 \geq 0$ such that, for all $(x_1, \ldots, x_m) \in \mathsf{E}^m$,

(24)
$$|\phi(x_1,\ldots,x_m)|_{\mathsf{F}} \le \lambda_2 (1+|x_1|_{\mathsf{E}}^{p+1}+\cdots+|x_m|_{\mathsf{E}}^{p+1}).$$

Denote by $|\phi|_{\text{Li}(p)}$ the smallest λ satisfying (23) and (24). Li(E, m, F; p) is called the linear space of p-weighted Lipschitz mappings and $|\phi|_{\text{Li}(p)}$ the

p-weighted Lipschitz norm of ϕ . We now state the exponential forgetting property. In the sequel, for any integrable r.v. Z, $\mathbb{E}^{\mathcal{F}}[Z]$ denotes the conditional expectation of Z given the σ -field \mathcal{F} and inequalities involving random variables will be meant in the *almost sure* sense.

Proposition 11. Let $p \geq 0$. Assume that \mathbf{U}_{p+1}^{\star} is finite and that (22) is satisfied for some C > 0 and $\rho < 1$. Let $\phi \in \mathrm{Li}(\mathbb{R}^d, m, \mathbb{R}; p)$, where m is a positive integer. For all $k \geq 0$ let us denote $\mathcal{F}_k := \sigma(\mathbf{Z}_0, \mathbf{U}_l, 1 \leq l \leq k)$. Then, there exist constants C_1 and C_2 (depending only on $C, \rho, \mathbf{U}_{p+1}^{\star}, B^{\star}$ and p) such that, for all $0 \leq l \leq k$ and $0 = j_1 < \cdots < j_m$,

(25)
$$\left| \mathbb{E}^{\mathcal{F}_k} \left[\mathbf{\Phi}_{k+r} \right] \right| \leq C_1 \, m \, |\phi|_{\operatorname{Li}(p)} \left(1 + \rho^{r(p+1)} \, |\mathbf{Z}_k|^{p+1} \right),$$

$$\left| \mathbb{E}^{\mathcal{F}_k} \left[\mathbf{\Phi}_{k+r} \right] - \mathbb{E}^{\mathcal{F}_l} \left[\mathbf{\Phi}_{k+r} \right] \right| \leq C_2 \, m \, \rho^r \, |\phi|_{\operatorname{Li}(p)} \left(1 + |\mathbf{Z}_k|^{p+1} + \mathbb{E}^{\mathcal{F}_l} \left[|\mathbf{Z}_k|^{p+1} \right] \right),$$

$$\left| \mathbb{E}^{\mathcal{F}_k} \left[\mathbf{\Phi}_{k+r} \right] - \mathbb{E} \left[\mathbf{\Phi}_{k+r} \right] \right| \leq C_2 \, m \, \rho^r \, |\phi|_{\operatorname{Li}(p)} \left(1 + |\mathbf{Z}_k|^{p+1} + \mathbb{E} \left[|\mathbf{Z}_k|^{p+1} \right] \right),$$

where
$$\Phi_j = \phi(\mathbf{Z}_{j+j_1}, \dots, \mathbf{Z}_{j+j_m})$$
 for all $j \geq 0$.

Proof. Define $H_{k,r}$ as the $\mathbb{R}^d \to \mathbb{R}$ function mapping $x \in \mathbb{R}^d$ to

$$\mathbb{E}\left[\phi(h_{k,r+j_1}(\mathbf{x},\mathbf{U}_{k+1},\ldots,\mathbf{U}_{k+r+j_1}),\ldots,h_{k,r+j_m}(\mathbf{x},\mathbf{U}_{k+1},\ldots,\mathbf{U}_{k+r+j_m}))\right],$$

where, for all $i, j \in \mathbb{N}$ and $(\mathbf{x}, \mathbf{u}_1, \dots, \mathbf{u}_j) \in \mathbb{R}^d \times \mathbb{R}^{qj}$,

$$h_{i,j}(\mathbf{x}, \mathbf{u}_1, \dots, \mathbf{u}_j) = \alpha(i+j, i) \mathbf{x} + \sum_{k=1}^{j} \alpha(i+j, i+k) B_{i+k} \mathbf{u}_k,$$
with $\alpha(l, l) = I$ and $\alpha(l+m, l) := \prod_{k=l+1}^{l+m} A_k, \quad l \ge 0, m \ge 1.$

We thus have, for all $k, r \geq 0$,

(26)
$$\mathbb{E}^{\mathcal{F}_k} \left[\mathbf{\Phi}_{k+r} \right] = H_{k,r}(\mathbf{Z}_k).$$

Using (22), observe that, for all $i, j \in \mathbb{N}$ and $(\mathbf{x}, \mathbf{y}, \mathbf{u}_1, \dots, \mathbf{u}_j) \in \mathbb{R}^{2d} \times \mathbb{R}^{qj}$,

$$|h_{i,j}(\mathbf{x},\mathbf{u}_1,\ldots,\mathbf{u}_j)-h_{i,j}(\mathbf{y},\mathbf{u}_1,\ldots,\mathbf{u}_j)|\leq C\,\rho^j\,|\mathbf{x}-\mathbf{y}|,$$

$$|h_{i,j}(\mathbf{x}, \mathbf{u}_1, \dots, \mathbf{u}_j)| \le C \left(\rho^j |\mathbf{x}| + \sum_{k=1}^j \rho^{j-k} B^* |\mathbf{u}_k|\right).$$

Using these bounds among with the Minkowski inequality, the definition of $H_{k,r}$ above and the assumptions on ϕ , we easily obtain, for all $(\mathbf{x}, \mathbf{y}) \in \mathbb{R}^{2d}$,

$$|H_{k,r}(\mathbf{x}) - H_{k,r}(\mathbf{y})| \le c_1 |\phi|_{\mathrm{Li}(p)} \left(\sum_{i=1}^m \rho^{r+j_i} \right) |\mathbf{x} - \mathbf{y}|$$

$$\left[1 + \sum_{i=1}^m \rho^{(r+j_i)p} (|\mathbf{x}|^p + |\mathbf{y}|^p) + 2 \sum_{i=1}^m \left(\sum_{k=1}^{r+j_i} \rho^{r+j_i-k} B^* \mathbf{U}_p^* \right)^p \right],$$

$$|H_{k,r}(\mathbf{x})| \le c_1 |\phi|_{\mathrm{Li}(p)}$$

$$\left[1 + \sum_{i=1}^m \rho^{(r+j_i)(p+1)} |\mathbf{x}|^{p+1} + \sum_{i=1}^m \left(\sum_{k=1}^{r+j_i} \rho^{r+j_i-k} B^* \mathbf{U}_{p+1}^* \right)^{p+1} \right],$$

where c_1 is a constant only depending on C, ρ and p. Observing that $\sum_{k=1}^{j} \rho^{j-k} \leq 1/(1-\rho)$ and $\sum_{i=1}^{m} \rho^{j_i\alpha} \leq 1/(1-\rho^{\alpha})$ for all $j \geq 1$, $0 = j_1 < \cdots < j_m$ and $\alpha > 0$, we obtain

(27)
$$|H_{k,r}(\mathbf{x})| \le C_1 |\phi|_{\mathrm{Li}(p)} m \left(1 + \rho^{r(p+1)} |\mathbf{x}|^{p+1}\right),$$

(28)
$$|H_{k,r}(\mathbf{x}) - H_{k,r}(\mathbf{y})| \le C_1 m \rho^r |\phi|_{\mathrm{Li}(p)} |x - y| (1 + \rho^{rp} (|\mathbf{x}|^p + |\mathbf{y}|^p)),$$

where C_1 is a constant only depending on B^* , \mathbf{U}_{p+1}^* , C, ρ and p. Eq. (25) follows from (26) and (27). Observe now that, for any probability measure ζ of \mathbb{R}^d ,

$$\left| H_{k,r}(\mathbf{x}) - \int H_{k,r}(\mathbf{y}) \, \zeta(d\mathbf{y}) \right| \le \int \left| H_{k,r}(\mathbf{x}) - H_{k,r}(\mathbf{y}) \right| \, \zeta(d\mathbf{y})$$

$$\le C_1 \, m \, \rho^r \, |\phi|_{\mathrm{Li}(p)} \, \int |\mathbf{x} - \mathbf{y}| \, (1 + |\mathbf{x}|^p + |\mathbf{y}|^p) \, \zeta(d\mathbf{y}),$$

where we used (28). The two last bounds of the proposition follow by using (26), $|\mathbf{x} - \mathbf{y}| \leq |\mathbf{x}| + |\mathbf{y}|$ and by choosing ζ respectively equal to the conditional distribution of \mathbf{Z}_k given $(\mathbf{Z}_0, \mathbf{U}_1, \dots, \mathbf{U}_l)$ and to the distribution of \mathbf{Z}_k .

4. Proof of Theorem 2

We denote by \mathbb{M}_d and \mathbb{M}_d^+ the space of $d \times d$ real matrices and the subspace of positive semi-definite symmetric matrices, respectively. For all $A \in \mathbb{M}_d$, we let $\lambda_{\min}(A)$, $\lambda_{\max}(A)$ and $|\lambda|_{\max}(A)$ denote the minimum eigenvalue, the maximum eigenvalue and the spectral radius of the matrix A, respectively, that is, $\lambda_{\min}(A) := \min(\operatorname{sp}(A))$, $\lambda_{\max}(A) := \max(\operatorname{sp}(A))$ and

 $|\lambda|_{\max}(A) := \max |\operatorname{sp}(A)|$, where $\operatorname{sp}(A)$ denotes the set of eigenvalues of A. The proof is derived in five steps and relies on intermediary results which will be repeatedly used throughout the paper.

Step 1: exponential Stability of the TVAR model. We have

Lemma 12. Let $0 < \rho_0 < \rho$ and L a positive constant. Then there exists C > 0 such that, for all $A \in \mathbb{M}_d$ with $|A| \leq L$ and $|\lambda|_{\max}(A) \leq \rho_0$, and for all $k \in \mathbb{N}$, $|A^k| \leq C\rho^k$.

Proof. We apply Theorem VII.1.10 of Dunford and Schwartz (1958). Let $\gamma = \{z \in \mathbb{C} : |z| = \rho\}$. Then, for any A such that $|\lambda|_{\max}(A) \leq \rho_0$ and for all $k \in \mathbb{N}$,

$$A^{k} = \frac{1}{2\pi i} \int_{\gamma} z^{k} (z I - A)^{-1} dz.$$

Hence, putting $z = \rho e^{iy}$, $y \in (-\pi, \pi)$, we have

$$|A^k| \le \frac{\rho^k}{2\pi} \int_{-\pi}^{\pi} |(I - A/(\rho e^{iy}))^{-1}| \, dy.$$

Let $G := \{A \in \mathbb{M}_d : |A| \leq L, |\lambda|_{\max}(A) \leq \rho_0\}$. G is a compact set and $(z,A) \mapsto |(I-A/z)^{-1}|$ is continuous over $\gamma \times G$ so that it is uniformly bounded. The proof follows.

For all $t \in [0,1]$ and for all $\boldsymbol{\theta} = (\theta_1, \dots, \theta_d) : [0,1] \to \mathbb{R}^d$, let $\Theta(t, \boldsymbol{\theta})$ denote the companion matrix defined by

(29)
$$\Theta(t, \boldsymbol{\theta}) = \begin{bmatrix} \theta_1(t) & \theta_2(t) & \dots & \theta_d(t) \\ 1 & 0 & \dots & 0 \\ 0 & 1 & 0 & \dots & 0 \\ \vdots & & & \vdots \\ 0 & \dots & 0 & 1 & 0 \end{bmatrix},$$

whose eigenvalues are the reciprocals of the roots of the autoregressive polynomial. Using this notation (2) can be rewritten as

(30)
$$\mathbf{X}_{k+1,n} = \Theta(k/n, \boldsymbol{\theta}) \mathbf{X}_{k,n} + \boldsymbol{\sigma}_{k+1,n} \epsilon_{k+1,n}, \quad 0 \le k \le n-1,$$

where $\boldsymbol{\sigma}_{k,n} = [\sigma(k/n) \ 0 \ \dots \ 0]^T$.

Proposition 13. Let $\beta \in (0,1]$, L > 0 and $0 < \rho < \tau < 1$. Then, there exists a constant M > 0 such that, for all $\theta \in \Lambda_d(\beta, L) \cap S(\rho)$ and $0 \le k < k + m \le n$,

(31)
$$\left| \prod_{l=k+1}^{k+m} \Theta(l/n, \boldsymbol{\theta}) \right| \le M \tau^m.$$

Proof. For notational convenience we use Λ and S as shorthand notations for $\Lambda_d(\beta, L)$ and $S(\rho)$. First note that $\Theta^* = \sup_{\boldsymbol{\theta} \in \Lambda \cap S} \sup_{t \in [0,1]} |\Theta(t, \boldsymbol{\theta})|$ is finite.

For any square matrices A_1, \ldots, A_r , we have

(32)

$$\prod_{k=1}^{r} A_k = A_1^r + (A_r - A_1)A_1^{r-1} + A_r(A_{r-1} - A_1)A_1^{r-2} + \dots + \prod_{k=3}^{r} A_k(A_2 - A_1)A_1.$$

By applying this decomposition to

(33)

$$\beta_n(k, i; \boldsymbol{\theta}) := \prod_{j=i+1}^k \Theta(j/n, \boldsymbol{\theta}), \quad 0 \le i < k \le n, \quad \beta_n(i, i; \boldsymbol{\theta}) = I, \quad 0 \le i \le n,$$

we have, for all $0 \le k < k + q \le n$,

$$|\beta_n(k+q,k;\boldsymbol{\theta})| \le |\Theta((k+1)/n,\boldsymbol{\theta})^q| + q \Theta^{\star q-1} \max_{1 \le j \le q} |\boldsymbol{\theta}((k+j)/n) - \boldsymbol{\theta}((k+1)/n)|.$$

Let us set $\tilde{\rho} \in (\rho, \tau)$. Hence, for all $\boldsymbol{\theta} \in \mathcal{S}(\rho)$ and $t \in [0, 1]$ we have $|\lambda|_{\max}(\Theta(t, \boldsymbol{\theta})) \leq \rho$. Since $\Theta^* < \infty$, by Lemma 12, we obtain

$$\sup_{\boldsymbol{\theta} \in \Lambda \cap \mathcal{S}} \sup_{t \in [0,1]} |\Theta^q(t,\boldsymbol{\theta})| \le C \, \tilde{\rho}^q, \quad q \in \mathbb{N}.$$

Observe that, for all $n \geq 1$,

$$\sup_{\boldsymbol{\theta} \in \Lambda} \max_{1 \le j \le q} \max_{0 \le k \le n-q} |\boldsymbol{\theta}((k+j)/n) - \boldsymbol{\theta}((k+1)/n)| \le L(q/n)^{\beta}.$$

Pick q and then N large enough so that $C\tilde{\rho}^q \leq \tau^q/2$ and $q \Theta^{\star q-1} L(q/N)^{\beta} \leq \tau^q/2$. The three last displays then give, for all $n \geq N$,

$$\sup_{\boldsymbol{\theta} \in \Lambda \cap \mathcal{S}} \max_{0 \le k \le (n-q)} |\beta_n(k+q,k;\boldsymbol{\theta})| \le \tau^q.$$

Write m = sq + t, $0 \le t < q$. For all $n \ge N$, $l \in \{1, ..., n - m\}$, and for all $\theta \in \Lambda \cap S$,

$$|\beta_n(l+m,l-1;\boldsymbol{\theta})| \le \Theta^{\star t} \prod_{i=0}^{s-1} |\beta_n(l+(i+1)q,l+iq;\boldsymbol{\theta})| \le (1+\Theta^{\star}/\tau)^q \tau^m.$$

The proof follows.

From Proposition 13 and Proposition 10, we get that, under (A1),

(34)
$$\sup_{(\boldsymbol{\theta},\sigma)\in\mathcal{C}}\sup_{0\leq k\leq n}\|\mathbf{X}_{k,n}\|_{q,\boldsymbol{\theta},\sigma}<\infty.$$

Eq. (31) and (34) are referred to as uniform exponential stability and uniform L^q boundedness respectively. The bound (34) may be extended to conditional moments as follows.

Proposition 14. Assume (A1) with $q \ge 1$ and let $p \in [1, q]$. Let $\beta \in (0, 1]$, L > 0, $0 < \rho < \tau < 1$ and $0 < \sigma_{-} \le \sigma_{+}$. Then, there exists a constant M such that, for all $(\theta, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_{-}, \sigma_{+})$ and $1 \le k \le l \le n$,

(35)
$$\left| \mathbb{E}_{\boldsymbol{\theta}, \sigma}^{\mathcal{F}_{k,n}} \left[|\mathbf{X}_{l,n}|^p \right] \right| \leq M \left(1 + \tau^{(l-k)} |\mathbf{X}_{k,n}|^p \right),$$

where $\mathcal{F}_{k,n} = \sigma\left(\mathbf{X}_{0,n}, X_{j,n}, 1 \leq j \leq k\right)$.

Proof. Eq. (31) is satisfied by Proposition 10. Then under **(A1)**, we apply Proposition 11 with $\phi(\mathbf{x}) = |\mathbf{x}|^p$, since $\phi \in \text{Li}(\mathbb{R}^d, 1, \mathbb{R}; p-1)$. Eq. (35) follows from (25).

Step 2: error decomposition. When studying recursive algorithms of the form (3) it is convenient to rewrite the original recursion in terms of the error defined by $\delta_{k,n} := \widehat{\boldsymbol{\theta}}_{k,n}(\mu) - \boldsymbol{\theta}_{k,n}, \ 0 \le k \le n$. Let us denote, for all $\nu \ge 0$ and $\mathbf{x} \in \mathbb{R}^d$,

(36)
$$L_{\nu}(\mathbf{x}) := \frac{\mathbf{x}}{1 + \nu |\mathbf{x}|^2} \quad \text{and} \quad F_{\nu}(\mathbf{x}) := L_{\nu}(\mathbf{x})\mathbf{x}^T.$$

The tracking error process $\{\delta_{k,n}, 0 \leq k \leq n\}$ obeys the following sequence of linear stochastic difference equations. For all $0 \leq k < n$,

(37)
$$\delta_{k,n} = \delta_{k,n}^{(u)} + \delta_{k,n}^{(v)} + \delta_{k,n}^{(w)},$$

$$\delta_{k+1,n}^{(u)} := (I - \mu \operatorname{F}_{\mu}(\mathbf{X}_{k,n})) \, \delta_{k,n}^{(u)}, \quad \delta_{0,n}^{(u)} = -\boldsymbol{\theta}_{0},$$

$$\delta_{k+1,n}^{(v)} := (I - \mu \operatorname{F}_{\mu}(\mathbf{X}_{k,n})) \delta_{k,n}^{(v)} + \mu \operatorname{L}_{\mu}(\mathbf{X}_{k,n}) \, \sigma_{k+1,n} \epsilon_{k+1,n}, \quad \delta_{0,n}^{(v)} = 0,$$

$$\delta_{k+1,n}^{(w)} := (I - \mu \operatorname{F}_{\mu}(\mathbf{X}_{k,n})) \, \delta_{k,n}^{(w)} + (\boldsymbol{\theta}_{k,n} - \boldsymbol{\theta}_{k+1,n}), \quad \delta_{0,n}^{(w)} = 0.$$

 $\{\delta_{k,n}^{(\mathrm{u})}\}$ takes into account the way the successive estimates of the regression coefficients forget the initial error. Making a parallel with classical non-parametric function estimation, $\delta_{k,n}^{(\mathrm{w})}$ plays the role of a bias term (this term

cancels when the function $t \mapsto \boldsymbol{\theta}(t)$ is constant), whereas $\delta_{k,n}^{(v)}$ is a stochastic disturbance. It should be stressed that the "bias term" is non-deterministic as soon as $t \mapsto \boldsymbol{\theta}(t)$ is not constant. The transient term simply writes, for all $0 \le k < n$,

(38)
$$\delta_{k+1,n}^{(u)} = \Psi_n(k, -1; \mu) \ \delta_0^{(u)},$$

where, for all $\mu \geq 0$ and $-1 \leq j < k \leq n$,

(39)
$$\Psi_n(j,j;\mu) := I \text{ and } \Psi_n(k,j;\mu) = \prod_{l=j+1}^k (I - \mu \, \mathcal{F}_{\mu}(\mathbf{X}_{l,n})).$$

Let us finally define the following increment processes, for all $0 \le k < n$,

$$\xi_{k,n}^{(\mathrm{w})} := \boldsymbol{\theta}_{k,n} - \boldsymbol{\theta}_{k+1,n} \quad \text{and} \quad \xi_{k,n}^{(\mathrm{v})} := \mu \operatorname{L}_{\mu}(\mathbf{X}_{k,n}) \, \sigma_{k+1,n} \, \epsilon_{k+1,n}.$$

According to these definitions, $\{\delta_{k,n}^{(\mathrm{v})}\}_{1\leq k\leq n}$ and $\{\delta_{k,n}^{(\mathrm{w})}\}_{1\leq k\leq n}$ obey a generic sequence of inhomogeneous stochastic recurrence equations of the form

(40)
$$\delta_{k+1,n}^{(\bullet)} = (I - \mu \operatorname{F}_{\mu}(\mathbf{X}_{k,n})) \, \delta_{k,n}^{(\bullet)} + \xi_{k,n}^{(\bullet)} = \sum_{i=0}^{k} \Psi_n(k,i;\mu) \, \xi_{i,n}^{(\bullet)}, \quad 0 \le k < n.$$

In view of (38) and (40), it is clear that the stability of the product of random matrices $\Psi_n(k,i;\mu)$ plays an important role in the limiting behavior of the estimation error.

Step 3: stability of the recursive algorithm. The following stability result for the product $\Psi_n(k,j;\mu)$ defined in (39) is an essential step for deriving risk bounds for the estimator $\hat{\theta}_n$.

Theorem 15. Assume (A1) with $q \ge 4$. Let $\beta \in (0,1]$, L > 0, $0 < \rho < 1$, and $0 < \sigma_{-} \le \sigma_{+}$. Then for all $p \ge 1$, there exist constants $M, \delta > 0$ and $\mu_{0} > 0$, such that, for all $0 \le j \le k \le n$, $\mu \in [0, \mu_{0}]$ and $(\theta, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_{-}, \sigma_{+})$,

(41)
$$\|\Psi_n(k,j;\mu)\|_{p,\boldsymbol{\theta},\sigma} \le M(1-\delta\mu)^{k-j}.$$

Similar stability results have been obtained in the framework of classical recursive estimation algorithms (see for instance Guo (1994) and Priouret and Veretennikov (1995)) but cannot be applied directly to our non-stationary and non-parametric context. Let us sketch the main arguments of the proof

in a more general context. Let $\{A_k(\nu), k \geq 0, \nu > 0\}$ be a \mathbb{M}_d^+ -valued process such that

(C-1) for all $k \in \mathbb{N}$ and $\nu \in [0, \nu_1], |\nu A_k(\nu)| \leq 1$.

Here $A_0(\nu)$, $A_1(\nu)$,... corresponds to the matrices $F_{\nu}(\mathbf{X}_{j+1,n})$, $F_{\nu}(\mathbf{X}_{j+2,n})$,... which appear in the product $\Psi_n(k,j;\nu)$ for some fixed j, k and n. Taking p=1 in (41) (this can be done without loss of generality under (C-1)), we want to prove that $\mathbb{E}\left|\prod_{k=1}^{l}(I-\nu A_k(\nu))\right| \leq C(1-\delta\nu)^l$ for some positive constants C and δ .

First observe that it is sufficient to have $\mathbb{E}^{\mathcal{F}_k}[|I - \nu A_{k+1}(\nu)|] \leq 1 - \delta \nu$ or, more generally, for some fixed integer r and $\alpha > 0$,

(42)
$$\mathbb{E}^{\mathcal{F}_k} \left[\left| \prod_{l=k+1}^{k+r} (I - \nu A_l(\nu)) \right| \right] \le 1 - \alpha \nu.$$

To obtain this inequality, we expand the product and then use Lemma 31 so that

$$\left| \prod_{l=k+1}^{k+r} (I - \nu A_l(\nu)) \right| \le 1 - \nu \lambda_{\min} \left(\sum_{l=k+1}^{k+r} A_l(\nu) \right) + \mathcal{R},$$

where \mathcal{R} is some remainder term which will be controlled by using $\sum_{l=k+1}^{k+r} |A_l(\nu)|^s$ for some s>1. It turns out that, in the previous display, the conditional expectation of the RHS given \mathcal{F}_k may only be controlled on a set $\{\phi_k \leq R_1\}$ where $R_1 > 0$ and $\{\phi_k\}$ is a positive adapted sequence (in the TVAR context, ϕ_k corresponds to $|\mathbf{X}_{k,n}|$). This yields the two following conditions

(C-2) there exists $\alpha_1 > 0$ such that, for all $k \in \mathbb{N}$ and $\nu \in [0, \nu_1]$,

$$\mathbb{E}^{\mathcal{F}_k} \left[\lambda_{\min} \left(\sum_{l=k+1}^{k+r} A_l(\nu) \right) \right] \ge \alpha_1 \mathbf{I}(\phi_k \le R_1),$$

(C-3) there exists s > 1 and $C_1 > 0$ such that, for all $k \in \mathbb{N}$ and $\nu \in [0, \nu_1]$,

$$\mathbf{I}(\phi_k \le R_1) \sum_{l=k+1}^{k+r} \mathbb{E}^{\mathcal{F}_k} \left[|A_l(\nu)|^s \right] \le C_1.$$

In turn the inequality (42) may only be shown on the set $\{\phi_k \leq R_1\}$ and it remains to check that this happens for sufficiently many k's. For this we use a classical Lyapounov condition, namely

(C-4) there exist $\lambda < 1$, $B \ge 1$ and an adapted process $\{V_k, k \ge 0\}$ on $[1, \infty)$ such that, for all $k \in \mathbb{N}$,

$$\mathbb{E}^{\mathcal{F}_k} \left[V_{k+r} \right] \le \lambda \, V_k \, \mathbf{I}(\phi_k > R_1) + B \, V_k \, \mathbf{I}(\phi_k \le R_1).$$

The previous arguments yield the following general result, whose precise technical proof is postponed to Appendix A for convenience.

Theorem 16. Let $(\Omega, \mathcal{F}, \mathbb{P}, \{\mathcal{F}_k : k \in \mathbb{N}\})$ be a filtered space. Let $\{\phi_k, k \geq 0\}$ be a non-negative adapted process and, for any $\nu \geq 0$, let $A(\nu) := \{A_k(\nu), k \geq 0\}$ be an adapted \mathbb{M}_d^+ -valued process. Let $r \geq 1$, $R_1 > 0$ and $\nu_1 > 0$ such that (C-1)-(C-4) hold. Then, for any $p \geq 1$, there exist $C_0 > 0$, $\delta_0 > 0$ and $\nu_0 > 0$ only depending on $p, \nu_1, \alpha_1, R_1, C_1, s, B$ and λ such that, for all $\nu \in [0, \nu_0]$ and $n \geq 1$,

$$\mathbb{E}^{\mathcal{F}_0} \left[\left| \prod_{i=1}^n (I - \nu A_i(\nu)) \right|^p \right] \le C_0 e^{-\delta_0 \nu n} V_0.$$

Having this general result in hand, we now prove Theorem 15.

Proof of Theorem 15. To verify (C-2), which is referred to as the persistence of excitation property in control theory literature, we need some intermediary results which hold under the assumptions of Theorem 15.

Lemma 17. There exists C > 0 such that, for all $d \leq j \leq n$ and $(\theta, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_-, \sigma_+)$,

(43)
$$\lambda_{\min} \left(\mathbb{E}_{\boldsymbol{\theta}, \sigma} \left[F_{\nu}(\mathbf{X}_{j,n}) \right] \right) \ge C, \quad \nu \in [0, 1],$$

and, for all $0 \le k \le j - d$,

(44)
$$\lambda_{\min} \left(\mathbb{E}_{\boldsymbol{\theta}, \sigma}^{\mathcal{F}_{k,n}} \left[F_0(\mathbf{X}_{j,n}) \right] \right) \ge C,$$

(45)
$$\lambda_{\min} \left(\mathbb{E}_{\boldsymbol{\theta}, \sigma}^{\mathcal{F}_{k,n}} \left[F_{\nu}(\mathbf{X}_{j,n}) \right] \right) \ge \frac{C}{1 + |\mathbf{X}_{k,n}|^4}, \quad \nu \in [0, 1].$$

Proof. Let $\mathbf{x} \in \mathbb{R}^d$, $|\mathbf{x}| = 1$, $d \leq j \leq n$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C} := \mathcal{C}(\beta, L, \rho, \sigma_-, \sigma_+)$. Write

$$\mathbf{X}_{j,n} = \beta_n(j, j - d; \boldsymbol{\theta}) \, \mathbf{X}_{j-d,n} + \sum_{i=1}^d \beta_n(j, j - d + i; \boldsymbol{\theta}) \, \boldsymbol{\sigma}_{j-d+i,n} \, \epsilon_{j-d+i,n},$$

where β_n is defined by (33). We have $\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j-d,n}}\left[(\mathbf{x}^T\mathbf{X}_{j,n})^2\right] \geq \sigma_-^2 \mathbf{x}^T H_{j,n,d}(\boldsymbol{\theta}) \mathbf{x}$, where $H_{j,n,d}(\boldsymbol{\theta}) := C_{j,n,d}(\boldsymbol{\theta})C_{j,n,d}^T(\boldsymbol{\theta})$ is the controllability Gramian (see Kailath (1980))

$$C_{j,n,d}(\boldsymbol{\theta}) := [\beta_n(j,j-d+1;\boldsymbol{\theta})G \dots \beta_n(j,j;\boldsymbol{\theta})G], \text{ where } G := [1 \ 0 \ \dots \ 0]^T.$$

One easily shows that, for $i=1,\ldots,d,$ $\beta_n(j,j+1-i;\boldsymbol{\theta})G$ has a unit i-th coordinate and zero coordinates below. Hence $\det(H_{j,n,d}(\boldsymbol{\theta})) = \det(C_{j,n,d}(\boldsymbol{\theta})) = 1$. In addition, from exponential stability, we have, for all $(\boldsymbol{\theta},\sigma) \in \mathcal{C}$, $|C_{j,n,d}(\boldsymbol{\theta})| \leq M$ for some positive M not depending on (j,n). Hence, for all $d \leq j \leq n$ and $(\boldsymbol{\theta},\sigma) \in \mathcal{C}$,

$$\lambda_{\min}(H_{j,n,d}(\boldsymbol{\theta})) \ge \frac{\det(H_{j,n,d}(\boldsymbol{\theta}))}{\lambda_{\max}^{d-1}(H_{j,n,d}(\boldsymbol{\theta}))} \ge M^{-(d-1)}.$$

It follows that, for all $0 \le k \le j-d$, $\mathbf{x} \in \mathbb{R}^d$ such that $|\mathbf{x}| = 1$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$

(46)
$$\mathbf{x}^T \mathbb{E}_{\boldsymbol{\theta}, \sigma}^{\mathcal{F}_{k,n}} \left[\mathbb{F}_0(\mathbf{X}_{j,n}) \right] \mathbf{x} = \mathbb{E}_{\boldsymbol{\theta}, \sigma}^{\mathcal{F}_{k,n}} \left[\mathbb{E}_{\boldsymbol{\theta}, \sigma}^{\mathcal{F}_{j-d,n}} \left[(\mathbf{x}^T \mathbf{X}_{j,n})^2 \right] \right] \ge \sigma_-^2 M^{-(d-1)},$$

showing (44) for any $C \ge \sigma_-^2 M^{-(d-1)}$. Eq.(43) also follows for $\nu = 0$. For all $0 \le k \le j - d$ and for all $\nu \in [0, 1]$ write

$$\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{k,n}}\left[(\mathbf{x}^T\mathbf{X}_{j,n})^2\right] = \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{k,n}}\left[\frac{|\mathbf{x}^T\mathbf{X}_{j,n}|}{(1+\nu|\mathbf{X}_{j,n}|^2)^{1/2}}\left\{|\mathbf{x}^T\mathbf{X}_{j,n}|\left(1+\nu|\mathbf{X}_{j,n}|^2\right)^{1/2}\right\}\right].$$

The Cauchy–Schwarz inequality shows that

$$\mathbf{x}^{T} \mathbb{E}_{\boldsymbol{\theta}, \sigma}^{\mathcal{F}_{k, n}} \left[\mathbb{F}_{\nu}(\mathbf{X}_{j, n}) \right] \mathbf{x} \geq \frac{\left(\mathbb{E}_{\boldsymbol{\theta}, \sigma}^{\mathcal{F}_{k, n}} \left[(\mathbf{x}^{T} \mathbf{X}_{j, n})^{2} \right] \right)^{2}}{\mathbb{E}_{\boldsymbol{\theta}, \sigma}^{\mathcal{F}_{k, n}} \left[(\mathbf{x}^{T} \mathbf{X}_{j, n})^{2} (1 + \nu \, |\mathbf{X}_{j, n}|^{2}) \right]}.$$

Since $|\mathbf{x}^T \mathbf{X}_{j,n}| \leq |\mathbf{X}_{j,n}|$ for $|\mathbf{x}| = 1$, by applying (46), we get, for all $\nu \in [0,1]$,

$$\lambda_{\min}\left(\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{k,n}}\left[\mathcal{F}_{\nu}(\mathbf{X}_{j,n})\right]\right) \geq \frac{\left(\sigma_{-}^{2} \, M^{-(d-1)}\right)^{2}}{\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{k,n}}\left[|\mathbf{X}_{j,n}|^{2}\right] + \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{k,n}}\left[|\mathbf{X}_{j,n}|^{4}\right]}.$$

The proof of (45) then follows from (35). The proof of (43) is along the same lines.

Lemma 18. Let $\phi \in \text{Li}(\mathbb{R}^d, 1, \mathbb{R}; 1)$. There exists M > 0 such that for all $1 \le i \le k \le n, \ \nu \ge 0$ and $(\theta, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_-, \sigma_+)$,

$$\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[\left| \sum_{j=i+1}^{k} \left(\phi(\mathbf{X}_{j,n}) - \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[\phi(\mathbf{X}_{j,n}) \right] \right) \right|^{2} \right] \leq M \left(k - i \right) \left| \phi \right|_{\mathrm{Li}(1)}^{2} \left(1 + |\mathbf{X}_{i,n}|^{4} \right).$$

Proof. For $j \in \{i+1,\ldots,k\}$, denote $\Delta_j = \phi(\mathbf{X}_{j,n}) - \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} [\phi(\mathbf{X}_{j,n})]$. For all $i \leq j \leq l \leq n$, $\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} [\Delta_l] = \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} [\phi(\mathbf{X}_{l,n})] - \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} [\phi(\mathbf{X}_{l,n})]$. The model is uniformly exponentially and L^q stable (see Proposition 13 and (34)). From Proposition 11, there exists $\tau \in (\rho, 1)$ and $C_1, C_2 > 0$ such that for all $0 \leq i \leq j \leq l \leq n$, $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_-, \sigma_+)$,

$$\left| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} \left[\Delta_l \right] \right| \leq C_1 \, \tau^{l-j} \, |\phi|_{\mathrm{Li}(1)} \left(1 + |\mathbf{X}_{j,n}|^2 + |\mathbf{X}_{i,n}|^2 \right),$$

$$\left| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[\Delta_l^2 \right] \right| \leq C_2 \, |\phi|_{\mathrm{Li}(1)}^2 \left(1 + |\mathbf{X}_{i,n}|^4 \right).$$

From Proposition 14, we get that there exists $C_3 > 0$ such that for all $0 \le i \le j \le l \le n$, $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_-, \sigma_+)$,

$$\left| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[\Delta_{j} \Delta_{l} \right] \right| = \left| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[\Delta_{j} \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} \left[\Delta_{l} \right] \right] \right| \leq \left(\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[\Delta_{j}^{2} \right] \right)^{1/2} \left(\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[\left(\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} \left[\Delta_{l} \right] \right)^{2} \right] \right)^{1/2} \leq C_{3} \left| \phi \right|_{\mathrm{Li}(1)}^{2} \tau^{l-j} \left(1 + |\mathbf{X}_{i,n}|^{4} \right),$$

and the result follows.

Lemma 19. For all $R_1, \alpha_1 > 0$, there exists $r_0 \geq 1$ such that, for all $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_-, \sigma_+), r \geq r_0, n \geq r, k = 0, \dots, n-r$ and $\nu \in (0, 1]$,

(47)
$$\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{k,n}} \left[\lambda_{\min} \left(\sum_{l=k+1}^{k+r} F_{\nu}(\mathbf{X}_{l,n}) \right) \right] \ge \alpha_1 \mathbf{I}(|\mathbf{X}_{k,n}| \le R_1),$$

where F_{ν} is defined in (36). In addition, there exist constants $\delta > 0$ and $\mu_0 > 0$, such that, for all $(\theta, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_-, \sigma_+)$, $d \leq k \leq n$ and $\nu \in [0, \nu_0]$,

(48)
$$|I - \nu \mathbb{E}_{\boldsymbol{\theta}, \sigma}[F_{\nu}(\mathbf{X}_{k,n})]| \le 1 - \delta \nu.$$

Proof. For any symmetric matrix A, we have $|\lambda_{\min}(A)| \leq |\lambda|_{\max}(A) = |A|$ (recall that $|\cdot|$ denotes the operator norm) and $\lambda_{\min}(A) = \inf_{|\mathbf{x}|=1} x^T A x$. From the last assertion, it follows that for any symmetric matrix B having same size as A, $\lambda_{\min}(A+B) \geq \lambda_{\min}(A) + \lambda_{\min}(B)$. Therefore,

$$\lambda_{\min}(A) \ge \lambda_{\min}(B) + \lambda_{\min}(A - B) \ge \lambda_{\min}(B) - |A - B|$$
.

Applying these elementary facts, we get, for all $0 \le k < k + r \le n$,

$$\lambda_{\min} \left(\sum_{j=k+1}^{k+r} F_{\nu}(\mathbf{X}_{j,n}) \right) \geq \sum_{j=k+1}^{k+r} \lambda_{\min} \left(\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{k,n}} \left[F_{\nu}(\mathbf{X}_{j,n}) \right] \right) - \left| \sum_{j=k+1}^{k+r} \left(F_{\nu}(\mathbf{X}_{j,n}) - \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{k,n}} \left[F_{\nu}(\mathbf{X}_{j,n}) \right] \right) \right|.$$

From its definition in (36), $F_{\nu}(\mathbf{x}) \in \mathbb{M}_d^+$ for all $\mathbf{x} \in \mathbb{R}^d$ and $\nu \geq 0$, and

$$\sup_{\nu>0} |F_{\nu}|_{\mathrm{Li}(1)} < \infty$$

Applying Lemma 17 and Lemma 18, we obtain that there exist C, M > 0 such that, for all $(\theta, \sigma) \in \mathcal{C}$ and $0 \le k < k + r \le n$,

$$\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{k,n}} \left[\lambda_{\min} \left(\sum_{j=k+1}^{k+r} F_{\nu}(\mathbf{X}_{j,n}) \right) \right] \ge \frac{C \left(r - d \right)}{1 + |\mathbf{X}_{k,n}|^4} - M \sqrt{r} \left(1 + |\mathbf{X}_{k,n}|^4 \right)^{1/2}.$$

Now pick two positive numbers R_1 and α_1 . If $|\mathbf{X}_{k,n}| > R_1$, Eq. (47) is clearly satisfied. If $|\mathbf{X}_{k,n}| \leq R_1$, the last equation implies that, for all $0 \leq k < k + r \leq n$,

$$\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{k,n}} \left[\lambda_{\min} \left(\sum_{j=k+1}^{k+r} F_{\nu}(\mathbf{X}_{j,n}) \right) \right] \ge \frac{C(r-d)}{1+R_1} - M\sqrt{r} (1+R_1^4)^{1/2}.$$

We may thus find r_0 such that the RHS of this inequality is larger than or equal to α_1 for all $r \geq r_0$. This concludes the proof of (47).

From the uniform L^2 -boundedness and (49) we get that there exists M such that, for all $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$, $\nu \in [0, 1]$, $d \leq k \leq n$, $|\mathbb{E}_{\boldsymbol{\theta}, \sigma}[F_{\nu}(\mathbf{X}_{k,n})]| \leq M$. Thus, using Lemma 31, for all $\nu \in [0, 1/M]$, $d \leq k \leq n$, and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$,

$$|I - \nu \mathbb{E}_{\boldsymbol{\theta}, \sigma}[F_{\nu}(\mathbf{X}_{k,n})]| = 1 - \nu \lambda_{\min}(\mathbb{E}_{\boldsymbol{\theta}, \sigma}[F_{\nu}(\mathbf{X}_{k,n})]).$$

and the proof of (48) follows from Lemma 17.

We now turn back to the proof of Theorem 15 by applying Theorem 16 to the sequence

$$\{(A_l = F_{\nu}(\mathbf{X}_{j+1+l,n}), \mathcal{F}_{j+1+l,n}), l = 0, \dots, k-j-1\}$$

It thus remains to show that conditions (C-1)–(C-4) of Theorem 16 hold with constants $r, R_1, \nu_1, \alpha_1, C_1, \lambda, B$ and s which neither depend on $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$ nor

on j, k, n. Set $\alpha_1 = 1$ and, for l = 0, ..., k - j - 1, $V_l := 1 + |\mathbf{X}_{j+l,n}|$ and $\phi_l := |\mathbf{X}_{j+l,n}|$.

Condition (C-1). For all $\nu \geq 0$ and $\mathbf{x} \in \mathbb{R}^d$, $\nu |F_{\nu}(\mathbf{x})| = \nu |\mathbf{x}|/(1+\nu|\mathbf{x}|) \leq 1$, which yields (C-1).

Condition (C-2). From Lemma 19, we may choose r_0 only depending on R_1 and α_1 such that (C-2) holds for all $r \geq r_0$.

Condition (C-3). From Lemma 32, $\sup_{\nu \geq 0} ||F_{\nu}|^{q/2}|_{\text{Li}(q-1)} < \infty$. From Proposition 14 there exists M > 0 such that, for all $\nu \geq 0$, $0 \leq i < i + l \leq n$ and $(\theta, \sigma) \in \mathcal{C}$,

$$\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[|\mathcal{F}_{\nu}(\mathbf{X}_{i+l,n})|^{q/2} \right] \leq M \left(1 + \tau^{ql} |\mathbf{X}_{i,n}|^{q} \right).$$

Hence (C-3) is obtained with s = q/2 and $C_1 = Mr(1 + R_1^q)$.

Condition (C-4). Let $\tau \in (\rho, 1)$. From Proposition 14, there exists M such that, for all $0 \le i < i + r \le n$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$, $\mathbb{E}_{\boldsymbol{\theta}, \sigma}^{\mathcal{F}_{j,n}}[|\mathbf{X}_{i+r,n}|] \le M(1 + \tau^r |\mathbf{X}_{i,n}|)$; thus, for any $R_1 > 0$,

$$\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[1 + |\mathbf{X}_{i+r,n}| \right] \le M\tau^r |\mathbf{X}_{i,n}| + M + 1 \le (M\tau^r + \frac{M+1}{R_1} \mathbf{I}(|\mathbf{X}_{i,n}| > R_1)) (1 + |\mathbf{X}_{i,n}|) + (M+1) \mathbf{I}(|\mathbf{X}_{i,n}| \le R_1).$$

Choose $r \ge r_0$ and $R_1 > 0$ so that $M\tau^r + (M+1)/R_1 < 1$. Condition (C-4)) is then satisfied with $\lambda := M\tau^r + (M+1)/R_1 < 1$ and B = 1 + M.

Finally, we obtain, for some positive constants r, C_0, δ_0 and μ_0 , for all $\nu \in (0, \mu_0], (\boldsymbol{\theta}, \sigma) \in \mathcal{C}$ and $0 \le j < k \le n$ such that $n - j \ge r$,

$$\|\Psi_n(k,j;\nu)\|_{p,\boldsymbol{\theta},\sigma}^p = \mathbb{E}\left[\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j+1,n}}\left[\Psi_n(k,j;\nu)\right]\right] \leq C_0 e^{-\delta_0 \nu n} \left(1 + \|\mathbf{X}_{j+1,n}\|_{p,\boldsymbol{\theta},\sigma}\right).$$

The uniform boundedness (34) then yields (41) when $n-j \geq r$. The restriction $n-j \geq r$ above is needed because (C-4), (C-2) and (C-3) are well defined only for n-j < r. Now recall that (C-1) implies $|\Psi_n(k,j;\nu)| \leq 1$. The result for n-j < r (implying k-j < r) follows by taking $M \geq (1-\delta\nu_1)^{-r}$ in (41).

Step 4: error bounds. Similarly to (49), one easily shows that

$$\sup_{\nu>0} |L_{\nu}|_{Li(0)} < \infty,$$

where L is defined in (36). The L^q -boundedness of $\{\mathbf{X}_{k,n}, 1 \leq k \leq n\}$ then gives

(51)
$$\mathbf{F}_{q/2}^{\star} := \sup_{(\boldsymbol{\theta}, \sigma) \in \mathcal{C}} \sup_{\nu \geq 0} \sup_{0 \leq k \leq n} \| \mathbf{F}_{\nu}(\mathbf{X}_{k,n}) \|_{q/2, \boldsymbol{\theta}, \sigma} < \infty,$$

(52)
$$L_q^{\star} := \sup_{(\boldsymbol{\theta}, \sigma) \in \mathcal{C}} \sup_{\nu \geq 0} \sup_{0 \leq k \leq n} \| L_{\nu}(\mathbf{X}_{k,n}) \|_{q, \boldsymbol{\theta}, \sigma} < \infty.$$

From now on, for convenience, we let the same M, δ and μ_0 denote positive constants depending neither on the time indices i, j, k, n, \ldots , the step-size μ nor on the parameter (θ, σ) .

Applying (38) and (41), for all $u \ge 1$, there exists M > 0 and $\mu_0 > 0$ such that, for all $\mu \in (0, \mu_0]$, $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$ and $1 \le k \le n$,

(53)
$$\|\delta_{k,n}^{(\mathbf{u})}\|_{u,\boldsymbol{\theta},\sigma} \le M \left(1 - \delta\mu\right)^k |\boldsymbol{\theta}(0)|.$$

Define

$$\Xi_n^{(\bullet)}(k,k) := 0 \quad \text{and} \quad \Xi_n^{(\bullet)}(k,j) := \sum_{i=j}^{k-1} \xi_{i,n}^{(\bullet)}, \qquad 0 \le j < k \le n.$$

For all $1 \le j \le k \le n$, we have

$$\Psi_n(k-1,j;\mu) - \Psi_n(k-1,j-1;\mu) = \mu \Psi_n(k-1,j;\mu) F_\mu(\mathbf{X}_{j-1,n}).$$

By integration by parts, for all $1 \le k \le n$ and $\mu \in (0, \mu_0]$, (40) reads

(54)
$$\delta_{k,n}^{(\bullet)} = \Psi_n(k-1,0;\mu) \Xi_n^{(\bullet)}(k,0) + \mu \sum_{j=1}^{k-1} \Psi_n(k-1,j;\mu) F_{\mu}(\mathbf{X}_{j-1,n}) \Xi_n^{(\bullet)}(k,j).$$

By applying the Hölder inequality and using (41) and (51), we get that, for any $u \in (1, q/2)$, there exists M > 0 such that, for all $\mu \in [0, \mu_0]$, $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$ and $0 \le j \le k \le n$,

(55)
$$\|\Psi_n(k,j;\mu) F_{\mu}(\mathbf{X}_{j-1,n})\|_{u,\theta,\sigma} \le M (1 - \delta \mu)^{k-j+1}.$$

We consider now the two terms $\delta_{k,n}^{(\mathrm{w})}$ and $\delta_{k,n}^{(\mathrm{v})}$ separately. We have, for all $0 \le j < k \le n$,

(56)
$$\left|\Xi_n^{(\mathrm{w})}(k,j)\right| := \left|\sum_{i=j}^{k-1} \xi_{i,n}^{(\mathrm{w})}\right| = |\boldsymbol{\theta}_{j,n} - \boldsymbol{\theta}_{k,n}| \le |\boldsymbol{\theta}|_{\Lambda,\beta} n^{-\beta} (k-j)^{\beta}.$$

Inserting (56), (41) and (55) into (54), we thus get that there exist M > 0, $\delta > 0$ and $\mu_0 > 0$ such that, for all $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$, $1 \le k \le n$ and $\mu \in (0, \mu_0]$,

$$\|\delta_{k,n}^{(\mathbf{w})}\|_{u,\boldsymbol{\theta},\sigma} \leq M \left((1 - \delta \,\mu)^k \, (k/n)^{\beta} + \mu \sum_{j=1}^{k-1} (1 - \delta \,\mu)^{k-j} \, ((k-j)/n)^{\beta} \right).$$

By Lemma 33, picking $u \geq p$, we obtain, for all $\mu \in (0, \mu_0]$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$,

(57)
$$\|\delta_{k,n}^{(\mathbf{w})}\|_{p,\boldsymbol{\theta},\sigma} \leq M |\boldsymbol{\theta}|_{\Lambda,\beta} (\mu n)^{-\beta}, \quad 1 \leq k \leq n.$$

We finally bound $\delta^{(v)}$. Note that, for each n > 1, $\{\sigma_{i,n} L_{\mu}(\mathbf{X}_{i,n}) \epsilon_{i+1,n}, i = 1, \ldots, n-1\}$ is an $\mathcal{F}_{i+1,n}$ -adapted martingale increment sequence. The Burkhölder inequality (see (Hall and Heyde, 1980, Theorem 2.12)) and (52) give, for all $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$ and $\mu \geq 0$,

(58)
$$\left\| \sum_{i=j}^{k} \sigma_{i,n} L_{\mu}(\mathbf{X}_{i,n}) \, \epsilon_{i+1,n} \right\|_{q,\boldsymbol{\theta},\sigma} \leq M \, (k-j+1)^{1/2}, \quad 1 \leq j \leq k < n.$$

By Lemma 33, using (54), (58), and (55) with u such that 1/p = 1/u + 1/q, we get, for all $\mu \in (0, \mu_0]$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$,

(59)
$$\|\delta_{k,n}^{(v)}\|_{p,\boldsymbol{\theta},\sigma} \le M\sqrt{\mu}, \quad 1 \le k \le n.$$

Eq. (13) easily follows from (37), (53), (57) and (59) and Theorem 2 is obtained.

5. Proof of Theorem 4

By writing $\hat{\delta}_n := \mathbf{u}^T \hat{\boldsymbol{\delta}}_n$, (14) simply means that, for all real-valued estimator $\hat{\delta}_n(\mathbf{X}_{0,n}, X_{1,n}, \dots, X_{n,n})$ and $\mathbf{u} = (u_1, \dots, u_d) \in \mathbb{R}^d$ such that $|\mathbf{u}| = 1$,

$$\sup_{(\boldsymbol{\theta}, \boldsymbol{\sigma}) \in \mathcal{C}} \mathbb{E}_{\boldsymbol{\theta}, \boldsymbol{\sigma}} \left[\left| \widehat{\delta}_n - \mathbf{u}^T \boldsymbol{\theta}(t) \right|^2 \right] \ge \alpha \, n^{\frac{-2\beta}{1+2\beta}}.$$

Denote by $\lfloor \beta \rfloor$ the largest integer strictly smaller than β . Let $\phi : \mathbb{R} \to \mathbb{R}$ be a C^{∞} symmetric function decreasing on \mathbb{R}_+ such that $\phi(0) = 1$, $\phi(u) = 0$ for all $|u| \geq 1$ and

(60)
$$\sup_{x,y \in \mathbb{R}, x \neq y} \frac{|\phi^{(\lfloor \beta \rfloor)}(x) - \phi^{(\lfloor \beta \rfloor)}(y)}{|x - y|^{\beta - \lfloor \beta \rfloor}} \le 1.$$

Let $\lambda : \mathbb{R} \to \mathbb{R}_+$ be a C^1 p.d.f. with respect to the Lebesgue measure vanishing outside the segment [-1, +1] and such that

(61)
$$\int_{-1}^{1} \left(\frac{\dot{\lambda}(x)}{\lambda(x)} \right)^{2} \lambda(x) \, dx < \infty ,$$

where $\dot{\lambda}$ is the derivative of the density λ . Let (v_n) and (w_n) be two non-decreasing sequence of positive numbers to be specified later such that

(62)
$$\lim_{n \to \infty} \left(v_n^{-1} + w_n^{-1} + n^{-1} w_n \right) = 0 \quad \text{and} \quad \sup_{n > 0} v_n^{-1} w_n^{\beta} \le L.$$

Let $t \in [0,1]$ and $\mathbf{u} \in \mathbb{R}^d$ such that $|\mathbf{u}| = 1$. Define $\phi_{n,t} : [0,1] \to \mathbb{R}^d$, $s \mapsto \phi_{n,t}(s) := \phi((s-t)w_n)\mathbf{u}$ and let $\sigma : [0,1] \to \mathbb{R}^+$, $s \mapsto \sigma(s) = \sigma_+$. For $n \ge 1$, define

(63)
$$X_{k+1,n} = \eta \phi((k/n-t)w_n)\mathbf{u}^T \mathbf{X}_{k,n} + (\sigma_+) \epsilon_{k+1,n}, \quad 0 \le k \le n-1.$$

From (A2), it follows that, for all $0 \le k \le n-1$,

$$x \mapsto p_{k+1,n}((x - \eta \phi_{n,t}(k/n)\mathbf{u}^T\mathbf{y})/\sigma_+)/\sigma_+$$

is the conditional density of $X_{k+1,n}$ given $\mathbf{X}_{k,n} = \mathbf{y}$ and parameter η . Since the distribution of $\mathbf{X}_{0,n}$ does not depend on η and using the conditional densities above to compute the joint density of $\{\mathbf{X}_{0,n}, X_{1,n}, \ldots, X_{n,n}\}$, the Fisher information associated to the one-dimensional parametric model defined by (63) is

$$\mathcal{I}_n(\eta) = \sigma_+^{-2} \mathbb{E}_{\eta \phi_{n,t},\sigma} \left[\left(\sum_{k=1}^n \phi_{n,t}(k/n) \mathbf{u}^T \mathbf{X}_{k-1,n} \frac{\dot{p}_{k,n}}{p_{k,n}} (\epsilon_{k,n}) \right)^2 \right].$$

Now, under (A2), the summand in this equation is a martingale increments sequence whose variances are bounded by $\phi_{n,t}^2(k/n) \mathbb{E}_{\eta\phi_{n,t},\sigma}[(u^T\mathbf{X}_{k-1,n})^2] \mathcal{I}_{\epsilon}$. Since |u| = 1, $(u^T\mathbf{X})^2 \leq |\mathbf{X}|^2$, and we finally obtain

(64)
$$\mathcal{I}_n(\eta) \le \sigma_+^{-2} \mathcal{I}_{\epsilon} \sum_{k=1}^n \phi_{n,t}^2(k/n) \, \mathbb{E}_{\eta \phi_{n,t},\sigma}[|\mathbf{X}_{k-1,n}|^2].$$

From (60) and (62), for all $\eta \in [-v_n^{-1}, v_n^{-1}]$,

$$|\eta \phi_{n,t}|_{\Lambda,\beta} \leq \frac{w_n^{\lfloor \beta \rfloor}}{v_n} \sup_{0 \leq s \leq s' \leq 1} \frac{|\phi^{(\lfloor \beta \rfloor)}((s'-t)w_n) - \phi^{(\lfloor \beta \rfloor)}((s-t)w_n)|}{|s'-s|^{\beta-\lfloor \beta \rfloor}} \leq \frac{w_n^{\beta}}{v_n} \leq L,$$

and $|\eta\phi_{n,t}(0)| \leq v_n^{-1}$. Hence, for large enough n, $\eta\phi_{n,t} \in \Lambda_d(\beta, L)$, for all $\eta \in [-v_n^{-1}, v_n^{-1}]$. By construction, for $s \in [0, 1]$, the autoregressive polynomial of $\eta\phi_{n,t}$ is given by $1 - \eta\phi_{n,t}(s) \sum_{i=1}^d u_i z^i$. Since $\lim_{n\to\infty} v_n = \infty$, for any ρ , $0 < \rho < 1$, there exists N, such that, for all $n \geq N$, $\eta\phi_{n,t} \in \mathcal{S}(\rho)$ $\eta \in [-v_n^{-1}, v_n^{-1}]$, and thus $(\eta\phi_{n,t}, \sigma) \in \mathcal{C}$. Using (34) for bounding $\mathbb{E}_{\eta\phi_{n,t},\sigma}[|\mathbf{X}_{k-1,n}|^2]$ in (64), it follows that there exists M depending only on ρ , β , σ_+ and L such that, for all sufficiently large n and for all $\eta \in [-v_n^{-1}, v_n^{-1}]$,

$$\mathcal{I}_n(\eta) \le M \, \mathcal{I}_{\epsilon} \sum_{k \in \mathbb{Z}} \phi^2(kw_n/n - tw_n) .$$

Using that ϕ is C^1 and compactly supported, we have

$$\lim_{h \to 0} \sup_{x \in \mathbb{R}} h \left| \sum_{k \in \mathbb{Z}} \phi^2(kh - x) - \int \phi^2(t) \ dt \right| = 0.$$

Eq. (62) shows that, for large enough n and for all $\eta \in [-v_n^{-1}, v_n^{-1}]$ we have

(65)
$$\mathcal{I}_n(\eta) \le M \, \mathcal{I}_{\epsilon} \, n w_n^{-1} \int \phi^2(t) \, dt \, (1 + o(1)).$$

We get that, for all real valued estimator $\widehat{\delta}_n := \widehat{\delta}_n(\mathbf{X}_{0,n}, X_{1,n}, \dots, X_{n,n})$, as $n \to \infty$,

$$\sup_{(\boldsymbol{\theta},\sigma)\in\mathcal{C}} \mathbb{E}_{\boldsymbol{\theta},\sigma}[(\widehat{\delta}_n - \mathbf{u}^T \boldsymbol{\theta}(t))^2] \ge \int_{-v_n^{-1}}^{v_n^{-1}} v_n \lambda(v_n \eta) \, \mathbb{E}_{\eta\phi_{n,t},\sigma}[(\widehat{\delta}_n - \eta)^2] \, d\eta \ge$$

$$\left(\sup_{\eta \in [-v_n^{-1}, v_n^{-1}]} \mathcal{I}_n(\eta) + \mathcal{I}_n(\lambda)\right)^{-1} \ge \left(O(nw_n^{-1} + v_n^2)\right)^{-1},$$

where the first inequality is the Bayesian lower bound of the minimax risk (recall that, for n sufficiently large $(\eta\phi_{n,t},\sigma)\in\mathcal{C}$ for all $|\eta|\leq v_n^{-1}$), the second inequality is the so called van Trees inequality (see Gill and Levit (1995)) with $\mathcal{I}_n(\eta)$ denoting the Fisher Information of the translation model associated to the p.d.f. $v_n\lambda(v_n\cdot)$ and the last inequality is implied by (61) and (65). The proof is concluded by choosing $v_n=w_n^\beta$ and $w_n=n^{1/(1+2\beta)}$.

Remark 7. Theorem 4 easily extends to cases where the distribution of $\{\mathbf{X}_{0,n}, n \geq 1\}$ depends on $\boldsymbol{\theta}$ in a not too pathological way. Assume for instance that $\mathbf{X}_{0,n}$ follows the distribution of a stationary AR process with parameter $(\boldsymbol{\theta}(0), \sigma(0))$ and with a given white noise (see e.g. Dahlhaus

(1996b)). In this case, the lower bound (14) holds for t > 0 and n sufficiently large without further assumptions. This clearly follows from the proof: since the distribution of $\mathbf{X}_{0,n}$ only depends on $(\boldsymbol{\theta}(0), \sigma(0))$ and since, for t > 0 and n sufficiently large $(\eta \phi_{n,t}(0), \sigma(0)) = (0, \sigma_+)$ does not depend on η , the computation of $\mathcal{I}_n(\eta)$ applies and the proof goes on similarly.

6. Perturbation expansion of the error

In this section, we first derive several approximation results of the error terms $\delta^{(w)}$ and $\delta^{(v)}$ defined by (40) in Section 6.1. Computational estimates are then obtained in Section 6.3 and the proofs of Theorem 6 and Theorem 7 are finally obtained in Section 6.4 and Section 6.5 respectively.

6.1. **General approximation results.** Observe that Theorem 2 only provides a bound of the risk. The approach developed in this section relies upon a perturbation technique (see Aguech et al. (2000)). Decompose the LHS of (40) as $\delta_{k,n}^{(\bullet)} = J_{k,n}^{(\bullet,0)} + H_{k,n}^{(\bullet,0)}$, with $J_{0,n}^{(\bullet,0)} = 0$, $H_{0,n}^{(\bullet,0)} = 0$ and

$$J_{k+1,n}^{(\bullet,0)} = \left(I - \mu \operatorname{\mathbb{E}}_{\boldsymbol{\theta},\sigma}[\operatorname{F}_{\mu}(\mathbf{X}_{k,n})]\right) J_{k,n}^{(\bullet,0)} + \xi_{k,n}^{(\bullet)},$$

$$H_{k+1,n}^{(\bullet,0)} = \left(I - \mu \operatorname{F}_{\mu}(\mathbf{X}_{k,n})\right) H_{k,n}^{(\bullet,0)} + \mu \left(\operatorname{\mathbb{E}}_{\boldsymbol{\theta},\sigma}[\operatorname{F}_{\mu}(\mathbf{X}_{k,n})] - \operatorname{F}_{\mu}(\mathbf{X}_{k,n})\right) J_{k,n}^{(\bullet,0)}.$$

The inhomogeneous first-order difference equation satisfied by $J_{k,n}^{(\bullet,0)}$ yields

(66)
$$J_{k+1,n}^{(\bullet,0)} = \sum_{i=0}^{k} \psi_n(k,i;\mu,\boldsymbol{\theta},\sigma) \, \xi_{i,n}^{(\bullet)}, \quad 0 \le k < n,$$

where, for all $\mu \geq 0$, $0 \leq i < k \leq n$, and $(\boldsymbol{\theta}, \sigma)$,

$$\psi_n(i, i; \mu, \boldsymbol{\theta}, \sigma) := I \text{ and } \psi_n(k, i; \mu, \boldsymbol{\theta}, \sigma) := \prod_{j=i+1}^k \left(I - \mu \mathbb{E}_{\boldsymbol{\theta}, \sigma}[F_{\mu}(\mathbf{X}_{j,n}))\right).$$

For ease of notations, we write $\mathbf{F}_{k,n} := \mathbf{F}_{\mu}(\mathbf{X}_{k,n})$, $\overline{\mathbf{F}}_{k,n} := \mathbf{F}_{k,n}(\mu) - \mathbb{E}_{\boldsymbol{\theta},\sigma}[\mathbf{F}_{k,n}(\mu)]$, \mathbb{E} instead of $\mathbb{E}_{\boldsymbol{\theta},\sigma}$, $\Psi_n(k,i)$ instead of $\Psi_n(k,i;\mu)$, $\psi_n(k,i)$ instead of $\psi_n(k,i;\mu,\boldsymbol{\theta},\sigma)$ and so on. This decomposition of the error process $\{\delta_{k,n}^{(\bullet)}, 1 \leq k \leq n\}$ can be extended to further approximation order s > 0 as follows.

(67)
$$\delta_{k,n}^{(\bullet)} = J_{k,n}^{(\bullet,0)} + J_{k,n}^{(\bullet,1)} + \dots + J_{k,n}^{(\bullet,s)} + H_{k,n}^{(\bullet,s)}$$

where

$$J_{k+1,n}^{(\bullet,0)} = (I - \mu \mathbb{E}[\mathbf{F}_{k,n}]) J_{k,n}^{(\bullet,0)} + \xi_{k,n}^{(\bullet)}, \qquad J_{0,n}^{(\bullet,0)} = 0,$$

$$\vdots$$

$$J_{k+1,n}^{(\bullet,r)} = (I - \mu \mathbb{E}[\mathbf{F}_{k,n}]) J_{k,n}^{(\bullet,r)} + \mu \overline{\mathbf{F}}_{k,n} J_{k,n}^{(\bullet,r-1)}, \qquad J_{l,n}^{(\bullet,r)} = 0, \ 0 \le l < r,$$

$$\vdots$$

$$H_{k+1,n}^{(\bullet,s)} = (I - \mu \mathbf{F}_{k,n}) H_{k,n}^{(\bullet,s)} + \mu \overline{\mathbf{F}}_{k,n} J_{k,n}^{(\bullet,s)}, \qquad H_{l,n}^{(\bullet,s)} = 0, \ l = 0, \dots, s.$$

The processes $J_{k,n}^{(\bullet,r)}$ depend linearly on $\xi_{k,n}^{(\bullet)}$ and polynomially in the error $\overline{\mathbf{F}}_{k,n}$. We now show that $J^{(\mathbf{w},0)}$ and $J^{(\mathbf{v},0)}$, respectively defined by setting $\xi^{(\bullet)} = \xi^{(\mathbf{w})}$ and $\xi^{(\bullet)} = \xi^{(\mathbf{v})}$ in (66), are the main terms in the error terms $\delta^{(\mathbf{w})}$ and $\delta^{(\mathbf{v})}$ defined by (40).

Proposition 20. Assume (A1) with q > 4 and let $p \in [1, q/4)$. Let $\beta \in (0, 1]$, L > 0, $0 < \rho < 1$, and $0 < \sigma_{-} \leq \sigma_{+}$. Then, there exist constants M and $\mu_{0} > 0$, such that, for all $(\theta, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_{+}, \sigma_{-})$, $\mu \in (0, \mu_{0}]$ and $1 \leq k \leq n$,

(68)
$$|J_{kn}^{(w,0)}| \le M (n\mu)^{-\beta},$$

(69)
$$\|\delta_{k,n}^{(w)} - J_{k,n}^{(w,0)}\|_{p,\theta,\sigma} \le M \sqrt{\mu} (n\mu)^{-\beta}.$$

Proof. From (48) in Lemma 19 (which holds under the assumptions of Theorem 15, there exist $\delta > 0$, $\mu_0 > 0$ and M > 0 such that, for all $0 \le i \le k \le n$, $\mu \in [0, \mu_0]$ and $(\theta, \sigma) \in \mathcal{C}$,

(70)
$$|\psi_n(k, i; \mu, \boldsymbol{\theta}, \sigma)| \le M (1 - \delta \mu)^{k-i}.$$

Note that $\psi_n(k, i-1) - \psi_n(k, i) = -\mu \psi_n(k, i) \mathbb{E}[\mathbf{F}_{i,n}]$. As in (54), write, for all $1 \le k \le n$,

(71)
$$J_{k,n}^{(\mathbf{w},0)} = \psi_n(k-1,0) \,\Xi_n^{(\mathbf{w})}(k,0) + \mu \sum_{j=1}^{k-1} \psi_n(k-1,j) \,\mathbb{E}[\mathbf{F}_{j-1,n}] \,\Xi_n^{(\mathbf{w})}(k,j).$$

Using (56), (51) and Lemma 33 shows (68). By (67), we have $\delta^{(w)} - J^{(w,0)} = H^{(w,0)} = J^{(w,1)} + H^{(w,1)}$, where, for all $1 \le k \le n$,

(72)
$$J_{k,n}^{(w,1)} = \mu \sum_{j=0}^{k-1} \psi_n(k-1,j) \, \overline{\mathbf{F}}_{j,n} \, J_{j,n}^{(w,0)},$$

(73)
$$H_{k,n}^{(w,1)} = \mu \sum_{j=0}^{k-1} \Psi_n(k-1,j) \, \overline{\mathbf{F}}_{j,n} \, J_{j,n}^{(w,1)}.$$

Set $\phi_j(\mathbf{x}) = \psi_n(k-1, j) \mathbf{F}_{\mu}(\mathbf{x}) J_{j,n}^{(\mathbf{w},0)}, j = 0, \dots, k-1$. Note that, from (49), (70) and (68), for all $\mu \in (0, \mu_0], 0 \le j < k \le n$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$,

$$|\phi_j|_{\text{Li}(1)} \le |\psi_n(k-1,j)| |J_{j,n}^{(w,0)}| |F_\mu|_{\text{Li}(1)} \le M (1-\delta\mu)^{k-j} (\mu n)^{-\beta}.$$

By applying Proposition 29 componentwise, we get, for all $\mu \in (0, \mu_0]$, $1 \le k \le n$ and $(\theta, \sigma) \in \mathcal{C}$,

$$\left\| \sum_{j=0}^{k-1} \psi_n(k-1,j) \, \overline{\mathbf{F}}_{j,n} \, J_{j,n}^{(\mathbf{w},0)} \right\|_{q/2} \le M \, (\mu n)^{-\beta} \left(\sum_{j=0}^{k-1} (1 - \delta \, \mu)^{k-1-j} \right)^{1/2}.$$

Hence, for all $\mu \in (0, \mu_0]$, $1 \le k \le n$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$

(74)
$$\left\| J_{k,n}^{(\mathbf{w},1)} \right\|_{q/3} \le M \sqrt{\mu} (\mu n)^{-\beta}.$$

Let u be such that 2/q + 2/q + 1/u = 1/p. Thus, by Theorem 15 and (51), for all $\mu \in (0, \mu_0]$, $1 \le k \le n$ and $(\theta, \sigma) \in \mathcal{C}$,

(75)
$$||H_{k,n}^{(\mathbf{w},1)}||_{p} \leq \mu \sum_{j=1}^{k-1} ||\Psi_{n}(k-1,j)||_{u} ||\overline{\mathbf{F}}_{j,n}||_{q/2} ||J_{j,n}^{(\mathbf{w},1)}||_{q/2}$$

$$\leq M\mu^{3/2} (\mu n)^{-\beta} \sum_{j=1}^{k-1} (1-\delta\mu)^{k-j} \leq M\sqrt{\mu} (\mu n)^{-\beta}.$$

Proposition 21. Assume **(A1)** with $q \geq 7$ and let $p \in [1, 2q/11)$. Let $\beta \in (0,1]$, L > 0, $0 < \rho < 1$, and $0 < \sigma_{-} \leq \sigma_{+}$. Then, there exist constants M and μ_{0} such that for all $(\theta, \sigma) \in \mathcal{C}(\beta, L, \rho, \sigma_{-}, \sigma_{+})$, $\mu \in (0, \mu_{0}]$ and $1 \leq k \leq n$,

(76)
$$||J_{k,n}^{(\mathbf{v},0)}(\boldsymbol{\theta},\sigma)||_{q,\boldsymbol{\theta},\sigma} \le M\sqrt{\mu},$$

(77)
$$\|\delta_{k,n}^{(\mathbf{v})} - J_{k,n}^{(\mathbf{v},0)}(\boldsymbol{\theta},\sigma)\|_{p,\boldsymbol{\theta},\sigma} \le M \,\mu.$$

Proof. The Burkhölder inequality (see (Hall and Heyde, 1980, Theorem 2.12)) shows that for all $1 \le k \le n$, $\mu \in (0, \mu_0]$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$,

$$||J_{k,n}^{(v,0)}||_q \le M \,\mu \,\sigma_+ \,\mathcal{L}_q^{\star} \,\epsilon_q^{\star} \,\left(\sum_{j=0}^{k-1} |\psi_n(k-1,j)|^2\right)^{1/2} \le M \,\mu \,\left(\sum_{j=0}^{k-1} (1-\delta \,\mu)^2\right)^{1/2}$$

and (76) follows from (52) and (70). We now bound

$$J_{k,n}^{(v,1)} = \mu \sum_{j=0}^{k-1} \psi_n(k-1,j) \, \overline{\mathbf{F}}_{j,n} \, J_{j,n}^{(v,0)}, \quad 1 \le k \le n.$$

Let us pick $2 \le k \le n$. By plugging (66) with $\xi^{(\bullet)} = \xi^{(w)}$, we obtain

(78)
$$J_{k,n}^{(\mathbf{v},1)} = \mu^2 \sum_{0 \le i < j \le k-1} \phi_j(\mathbf{X}_{j,n}) \, \gamma_{i,j}(\mathbf{X}_{i,n}) \, \sigma_{i+1,n} \, \epsilon_{i+1,n}$$

where, for all $0 \le i < j \le k - 1$,

(79)
$$\phi_j(\mathbf{x}) := \psi_n(k-1,j) \, \mathbf{F}_{\mu}(\mathbf{x}) \quad \text{and} \quad \gamma_{i,j}(\mathbf{x}) := \psi_n(j-1,i) \, \mathbf{L}_{\mu}(\mathbf{x}).$$

From (49) and (50) we have, for all $0 \le i < j < k$, $\phi_j \in \text{Li}(\mathbb{R}^d, 1, \mathbb{R}^d \times \mathbb{R}^d; 1)$ and $\gamma_{i,j} \in \text{Li}(\mathbb{R}^d, 1, \mathbb{R}^d; 0)$ and, furthermore, from (70),

$$|\phi_j|_{\text{Li}(1)} \le M (1 - \delta \mu)^{k-j}$$
 and $|\gamma_{i,j}|_{\text{Li}(0)} \le M (1 - \delta \mu)^{j-i}$,

where, as usual, M and δ are positive constant neither depending on indices i, j, k, n, on $\mu \in [0, \mu_0]$ nor on $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$. The following uniform bounds follow.

$$\sup_{0 < j < t} |\phi_j|_{\mathrm{Li}(1)} \le M, \qquad \sum_{j=1}^{t-1} |\phi_j|_{\mathrm{Li}(1)} \le M \,\mu^{-1},$$

$$\sup_{0 \le i < j < t} |\gamma_{i,j}|_{\mathrm{Li}(0)} \le M, \qquad \sup_{0 < j < t} \left(\sum_{0 < i < j} |\gamma_{i,j}|_{\mathrm{Li}(0)}^2 \right)^{\frac{1}{2}} \le M \,\mu^{-1/2}.$$

By applying Proposition 30 componentwise, we obtain $\|J_{k,n}^{(v,1)}\|_{\frac{2q}{7}} \leq M \mu$ uniformly over $1 \leq k \leq n$, $\mu \in [0, \mu_0]$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$. As in (75), let u > 1 be such that $u^{-1} + 2/q + 7/2q = 1/p$. Then, for all $1 \leq k \leq n$, $\mu \in [0, \mu_0]$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$,

$$||H_{k,n}^{(\mathbf{v},1)}||_p \le \mu \sum_{j=1}^{k-1} ||\Psi_n(k-1,j)||_u ||\overline{\mathbf{F}}_{j,n}||_{\frac{q}{2}} ||J_{j,n}^{(\mathbf{v},1)}||_{\frac{2q}{7}},$$

and thus $\|H_{k,n}^{(\mathbf{v},1)}\|_p \leq M \,\mu^2 \,\mathbf{F}_{q/2}^{\star} \sum_{j=1}^{k-1} (1-\delta\mu)^{k-j}$ which yields (77).

6.2. **Proof of Proposition 5.** We may write, for all $1 \le k \le n$,

$$\mathbf{X}_{k,n} = \beta_n(k,0;\boldsymbol{\theta}) \, \mathbf{X}_{0,n} + \sum_{j=1}^k \beta_n(k,j;\boldsymbol{\theta}) \, \boldsymbol{\sigma}_{j,n} \, \epsilon_{j,n},$$

where σ and β_n are defined right after (30) and in (33), respectively. Thus,

$$\mathbb{E}_{\boldsymbol{\theta}, \boldsymbol{\sigma}}[\mathbf{X}_{k,n} \mathbf{X}_{k,n}^T] = \beta_n(k, 0) \, \mathbb{E}[\mathbf{X}_{0,n} \mathbf{X}_{0,n}^T] \, \beta_n(k, 0)^T + \\ \sum_{l=0}^{k-1} \beta_n(k, k-l) \, \boldsymbol{\sigma}_{k-l,n} \, (\beta_n(k, k-l) \, \boldsymbol{\sigma}_{k-l,n})^T.$$

Let $\theta \in \mathcal{S}(\rho)$ with $\rho > 1$, $t \in [0,1]$ and let $\{Z_k, k \in \mathbb{Z}\}$ denote the stationary $AR(\theta(t), \sigma(t))$ with i.i.d. centered unit variance innovations denoted by $\{\varepsilon_k\}_{k\in\mathbb{Z}}$. Recall that $\Sigma(t, \theta, \sigma)$ denotes the $d \times d$ covariance matrix of $\{Z_k, k \in \mathbb{Z}\}$. Then, using classical results on AR models (see Brockwell and Davis (1991)), we have $[Z_k \ldots Z_{k-d}]^T = \sum_{l\geq 0} \Theta^l(t, \theta) \sigma(t) \varepsilon_{k-l}$, where $\sigma(t) := [\sigma(t) \ 0 \ \ldots \ 0]^T$, Θ is defined by (29) and the convergence holds in the L^2 sense. It follows that

$$\Sigma(t, \boldsymbol{\theta}, \sigma) = \sum_{l=0}^{\infty} \Theta^l(t, \boldsymbol{\theta}) \, \boldsymbol{\sigma}(t) \, \left(\Theta^l(t, \boldsymbol{\theta}) \, \boldsymbol{\sigma}(t) \right)^T.$$

Denote $\Theta_{k,n} := \Theta(k/n, \boldsymbol{\theta})$ and $\Sigma_{k,n} := \Sigma(k/n, \boldsymbol{\theta}, \sigma)$. We obtain

$$(80) \quad \mathbb{E}_{\boldsymbol{\theta},\sigma}[\mathbf{X}_{k,n}\mathbf{X}_{k,n}^{T}] - \Sigma_{k,n} = \beta_{n}(k,0) \left(\mathbb{E}[\mathbf{X}_{0,n}\mathbf{X}_{0,n}^{T}] - \Sigma_{0,n} \right) \beta_{n}(k,0)^{T}$$

$$+ \sum_{l=0}^{k-1} \left(\beta_{n}(k,k-l) \, \boldsymbol{\sigma}_{k-l,n} \left(\beta_{n}(k,k-l) \, \boldsymbol{\sigma}_{k-l,n} \right)^{T} - \Theta_{k,n}^{l} \, \boldsymbol{\sigma}_{k,n} \left(\Theta_{k,n}^{l} \, \boldsymbol{\sigma}_{k,n} \right)^{T} \right)$$

$$+ \sum_{l=k}^{\infty} \left(\beta_{n}(k,0) \Theta_{0,n}^{l-k} \, \boldsymbol{\sigma}_{0,n} \left(\beta_{n}(k,0) \Theta_{0,n}^{l-k} \, \boldsymbol{\sigma}_{0,n} \right)^{T} - \Theta_{k,n}^{l} \, \boldsymbol{\sigma}_{k,n} \left(\Theta_{k,n}^{l} \, \boldsymbol{\sigma}_{k,n} \right)^{T} \right).$$

Note that, for any matrices A_1, \ldots, A_r and B_1, \ldots, B_r with compatible sizes,

(81)
$$\prod_{i=1}^{r} A_i - \prod_{i=1}^{r} B_i = \sum_{j=1}^{r} \left(\prod_{k=1}^{j-1} A_k \right) (A_j - B_j) \left(\prod_{k=j+1}^{r} B_k \right).$$

From uniform exponential stability (see Proposition 13 and its proof), there exists M such that for all $1 \le l \le k \le n$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}^*$, $|\beta_n(k, k-l)| \le M \tau^l$ and $|\Theta_{k,n}^l| \le M \tau^l$ and thus

$$|\beta_n(k, k-l) - \Theta_{k,n}^l| \le M \tau^l \sum_{j=0}^{l-1} |\Theta_{k-j,n} - \Theta_{k,n}| \le M n^{-\beta} \tau^l l^{\beta+1},$$

Similarly, there exists M such that, for all $1 \le l \le k \le n$ and $(\theta, \sigma) \in \mathcal{C}^*$,

$$\begin{aligned} |\Theta_{0,n}^{l-k} - \Theta_{k,n}^{l-k}| &\leq M \, n^{-\beta} \, \tau^{l-k} \, (l-k) \, k^{\beta}, \\ \left| \beta_{n}(k,k-l) \, \boldsymbol{\sigma}_{k-l,n} \, (\beta_{n}(k,k-l) \, \boldsymbol{\sigma}_{k-l,n})^{T} - \Theta_{k,n}^{l} \, \boldsymbol{\sigma}_{k,n} \, (\Theta_{k,n}^{l} \, \boldsymbol{\sigma}_{k,n})^{T} \right| \\ &\leq M \, n^{-\beta} \, \tau^{2l} \, l^{\beta+1}, \\ \left| \beta_{n}(k,0) \Theta_{0,n}^{l-k} \, \boldsymbol{\sigma}_{0,n} \, (\beta_{n}(k,0) \Theta_{0,n}^{l-k} \, \boldsymbol{\sigma}_{0,n})^{T} - \Theta_{k,n}^{l} \, \boldsymbol{\sigma}_{k,n} \, (\Theta_{k,n}^{l} \, \boldsymbol{\sigma}_{k,n})^{T} \right| \\ &\leq M \, n^{-\beta} \, \tau^{2l} \, l \, k^{\beta}. \end{aligned}$$

The result follows by inserting these bounds in (80).

6.3. Further approximation results. To derive tractable asymptotic risk estimates, we need to derive approximate expressions of $J_{k,n}^{(w,0)}$ and $J_{k,n}^{(v,0)}$. We first derive approximations of $\mathbb{E}[\mathbf{F}_{k,n}]$, $1 \le k \le n$ and related quantities.

Lemma 22. Assume (A1) with $q \ge 4$. Let $\beta \in (0,1]$, L > 0, $0 < \rho < \tau < 1$, and $0 < \sigma_{-} \le \sigma_{+}$. Then there exist positive constants δ, ν_{0} and M such that, for all $0 \le k \le l \le n$, $\nu \in [0, \nu_{0}]$ and $(\theta, \sigma) \in C^{*}(\beta, L, \rho, \sigma_{-}, \sigma_{+})$,

(82)
$$\left| (I - \nu \Sigma(l/n, \boldsymbol{\theta}, \sigma))^{l-k} \right| \le (1 - \delta \nu)^{l-k},$$

(83)
$$|\Sigma(l/n, \boldsymbol{\theta}, \sigma) - \mathbb{E}_{\boldsymbol{\theta}, \sigma}[F_{\nu}(\mathbf{X}_{k,n})]| \le M (\tau^k + n^{-\beta}(l-k+1)^{\beta} + \nu),$$

(84)
$$|\Sigma(l/n, \boldsymbol{\theta}, \sigma) - \mathbb{E}_{\boldsymbol{\theta}, \sigma}[L_{\nu}L_{\nu}^{T}(\mathbf{X}_{k,n})]| \leq M (\tau^{k} + n^{-\beta}(l-k+1)^{\beta} + \nu).$$

Proof. By continuity of $(\theta, z, t) \mapsto |\theta(z; t)|$ (see (6)) and since

$$\left\{\boldsymbol{\vartheta} \in \mathbb{R}^d : |\boldsymbol{\vartheta}| \le L, 1 - \sum_{i=1}^d \boldsymbol{\vartheta}_i z^i \ne 0 \text{ for all } |z| \le \rho^{-1}\right\}$$

is a compact set, there exist $\delta > 0$ and M > 0, such that, for all $(\theta, \sigma) \in \mathcal{C}^*$,

(85)
$$\delta \leq \inf_{|z|=1} \inf_{t \in [0,1]} \frac{\sigma^2(t)}{|\theta(z;t)|^2} \leq \lambda_{\min} \left(\Sigma(t; \boldsymbol{\theta}, \sigma) \right)$$
$$\leq \lambda_{\max} \left(\Sigma(t; \boldsymbol{\theta}, \sigma) \right) \leq \sup_{|z|=1} \sup_{t \in [0,1]} \frac{\sigma^2(t)}{|\theta(z;t)|^2} \leq M.$$

Eq. (82) then follows from Lemma 31. Similarly, there exists $M < \infty$, such that, for all $(\theta, \sigma) \in \mathcal{C}^*$ and all $0 \le s \le t \le 1$,

(86)
$$|\Sigma(t; \boldsymbol{\theta}, \sigma) - \Sigma(s; \boldsymbol{\theta}, \sigma)| \le M(t - s)^{\beta}.$$

By Proposition 5, we get, for all $1 \le k \le l \le n$ and $(\theta, \sigma) \in \mathcal{C}^*$,

(87)
$$|\Sigma(l/n; \boldsymbol{\theta}, \sigma) - \mathbb{E}[\mathbf{X}_{k,n}\mathbf{X}_{k,n}^T]| \le |\Sigma(l/n; \boldsymbol{\theta}, \sigma) - \Sigma(k/n; \boldsymbol{\theta}, \sigma)| + |\Sigma(k/n; \boldsymbol{\theta}, \sigma)) - \mathbb{E}[\mathbf{X}_{k,n}\mathbf{X}_{k,n}^T]| \le M (\tau^k + n^{-\beta} (l - k + 1)^{\beta}).$$

This is (83) and (84) with $\nu = 0$. One easily shows that, for all $\mathbf{x} \in \mathbb{R}^d$ and $\nu \geq 0$, $|F_{\nu}(\mathbf{x}) - \mathbf{x}\mathbf{x}^T| \leq \nu |\mathbf{x}|^4$ and $|L_{\nu}L_{\nu}^T(\mathbf{x})| \leq 2\nu |\mathbf{x}|^4$. Since $q \geq 4$, we deduce (83) and (84) for $\nu > 0$ from (87) and uniform L^4 boundedness. \square

Let $\rho \in (0,1)$ and let $\boldsymbol{\theta} \in \mathcal{S}(\rho)$ and $\sigma : [0,1] \to \mathbb{R}^+$. Define the following sequence of recurrence equations applying to some increment process $\{\xi_{k,n}^{(\bullet)}, 0 \leq k \leq n\}$.

(88)
$$\tilde{J}_{k+1,n}^{(\bullet)}(\boldsymbol{\theta},\sigma) := \sum_{j=0}^{k} (I - \mu \Sigma((k+1)/n,\boldsymbol{\theta},\sigma))^{k-j} \xi_{j,n}^{(\bullet)}, \quad 0 \le k \le n.$$

We now show that $J^{(\mathrm{w},0)}$ and $J^{(\mathrm{v},0)}$ may be approximated by $\tilde{J}^{(\mathrm{w})}$ and $\tilde{J}^{(\mathrm{v})}$ respectively defined by setting $\xi^{(\bullet)} = \xi^{(\mathrm{w})}$ and $\xi^{(\bullet)} = \xi^{(\mathrm{v})}$ in (88) and then we compute asymptotic equivalents of $\tilde{J}^{(\mathrm{w})}$ and of $\tilde{J}^{(\mathrm{v})}$'s variance respectively.

Proposition 23. Assume (A1) with $q \ge 4$. Let $\beta \in (0,1]$, L > 0, $\rho < 1$, and $0 < \sigma_{-} \le \sigma_{+}$. Then, there exist constants M > 0 and $\mu_{0} > 0$ such that, for all $(\theta, \sigma) \in C^{\star}(\beta, L, \rho, \sigma_{-}, \sigma_{+})$, $\mu \in (0, \mu_{0}]$ and $1 \le k \le n$,

(89)
$$|J_{k,n}^{(w,0)}(\boldsymbol{\theta},\sigma) - \tilde{J}_{k,n}^{(w)}(\boldsymbol{\theta},\sigma)| \le M (\mu n)^{-\beta} ((\mu n)^{-\beta} + \mu),$$

(90)
$$||J_{k,n}^{(\mathbf{v},0)}(\boldsymbol{\theta},\sigma) - \tilde{J}_{k,n}^{(\mathbf{v})}(\boldsymbol{\theta},\sigma)||_{q,\boldsymbol{\theta},\sigma} \le M\sqrt{\mu} ((\mu n)^{-\beta} + \mu).$$

Proof. Let $\Delta_n(k, j; \mu, \boldsymbol{\theta}, \sigma) := \psi_n(k, j; \mu, \boldsymbol{\theta}, \sigma) - (I - \mu \Sigma(k/n, \boldsymbol{\theta}, \sigma))^{k-1-j}$ for all $0 \le j \le k \le n$, $\mu \ge 0$ and $(\boldsymbol{\theta}, \sigma)$. The definitions of $J^{(\bullet,0)}$ and $\tilde{J}^{(\bullet)}$ yield

(91)
$$J_{k,n}^{(\bullet,0)} - \tilde{J}_{k,n}^{(\bullet)} = \sum_{j=0}^{k-1} \Delta_n(k-1,j) \xi_{j,n}^{(\bullet)}, \quad 1 \le k \le n.$$

From (81), we get, for all $0 \le j < k \le n, \, \mu \ge 0$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}^*$

$$\Delta_n(k-1,j) = \mu \sum_{i=j}^{k-1} \psi_n(k-1,i+1) \left(\mathbb{E}[\mathbf{F}_{i,n}] - \Sigma(k/n) \right) (I - \mu \Sigma(k/n))^{i-j-1}.$$

Using (70), (82) and (83), there exists $\delta > 0$, $\mu_0 > 0$ and M > 0, such that, for all $1 \le j < k \le n$, $\mu \in (0, \mu_0]$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}^*$

$$(92) |\Delta_n(k-1,j)| \le M \mu (1-\delta\mu)^{k-1-j} (\tau^j + n^{-\beta} (k-j)^{\beta+1} + \mu (k-j)).$$

We further write, for all $1 \le j < k \le n$,

$$\Delta_n(k-1,j) - \Delta_n(k-1,j-1) = \mu \, \Delta_n(k-1,j) \, \mathbb{E}[\mathbf{F}_{j-1,n}] + \mu \, (I - \mu \, \Sigma(k/n))^{k-j-1} \, (\mathbb{E}[\mathbf{F}_{j-1,n}] - \Sigma(k/n)).$$

Applying (92), (82) and (83) and observing that \mathbf{F}_1^{\star} is finite, we get that there exists $\delta > 0$, $\mu_0 > 0$, $\tau \in (\rho, 1)$ and M > 0, such that, for all $1 \leq j < k \leq n$, $\mu \in (0, \mu_0]$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}^{\star}$,

(93)
$$|\Delta_n(k-1,j) - \Delta_n(k-1,j-1)| \le M\mu(1-\delta\mu)^{k-j-1} \left(n^{-\beta} (k-j)^{\beta} (\mu(k-j)+1) + \mu (\mu(k-j)+1) + \tau^j\right).$$

By integrating (91) by parts, for all $1 \le k \le n$, $J_{k,n}^{(\bullet,0)} - \tilde{J}_{k,n}^{(\bullet)}$ reads

$$\Delta_n(k-1,0) \,\Xi_n^{(\bullet)}(k,0) + \sum_{j=1}^{k-1} (\Delta_n(k-1,j) - \Delta_n(k-1,j-1)) \,\Xi_n^{(\bullet)}(k,j).$$

Using (56) and (58) to bound $|\Xi_n^{(w)}|$ and $||\Xi^{(v)}|_n||_q$ respectively together with (92), (93) and Lemma 33, there exists $\delta > 0$, $\mu_0 > 0$ and M such that, for all $1 \le k \le n$, $\mu \in (0, \mu_0]$ and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$,

$$|J_{k,n}^{(\mathbf{w},0)} - \tilde{J}_{k,n}^{(\mathbf{w})}| \le M \left(\frac{(\mu \, n)^{-\beta} + \mu}{(\mu \, n)^{\beta}} + \mu \sum_{j=0}^{k-1} (1 - \delta \mu)^j \, \tau^{k-j} \, \left(\frac{j+1}{n} \right)^{\beta} \right),$$

$$||J_{k,n}^{(\mathbf{v},0)} - \tilde{J}_{k,n}^{(\mathbf{v})}||_q \le M \left(\sqrt{\mu} ((\mu \, n)^{-\beta} + \mu) + \mu^2 \sum_{j=0}^{k-1} (1 - \delta \mu)^j \, \tau^{k-j} \, \sqrt{j+1} \right).$$

Using Lemma 33, we have, for any $\alpha \geq 0$,

$$\sum_{j=0}^{k-1} (1-\delta\mu)^j \, \tau^{k-j} \, (j+1)^{\alpha} \le (1-\tau)^{-1} \, \sup_{j \in \mathbb{N}} (1-\delta\mu)^j \, (j+1)^{\alpha} \le C \, (1-\tau)^{-1} \, (\delta\mu)^{-\alpha}.$$

We thus obtain (89) and (90).

Proposition 24. Let $\beta \in (0,1]$, L > 0, $\rho < 1$, and $0 < \sigma_{-} \leq \sigma_{+}$ and let $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}^{\star}(\beta, L, \rho, \sigma_{-}, \sigma_{+})$. Let $t \in (0,1]$ and assume that there exists $\boldsymbol{\theta}_{t,\beta} \in \mathbb{R}^{d}$, L' > 0 and $\beta' > \beta$ such that (18) holds true for all $u \in [0,t]$. Then, for all $\mu \in (0, \mu_{0}]$ and $n \geq 1$,

$$\left|\tilde{J}_{[tn],n}^{(\mathrm{w})} - \frac{\Gamma(\beta+1)}{(\mu\,n)^{\beta}}\,\Sigma^{-\beta}(t)\,\boldsymbol{\theta}_{t,\beta}\right| \leq M\,\left((\mu\,n)^{-\beta'} + \frac{n^{-\beta} + (1-\delta\,\mu)^{tn}}{(\mu\,n)^{\beta}}\right),$$

where M and μ_0 are positive constants only depending on $\beta, L, \rho, \sigma_-, \sigma_+, L'$ and β' .

Proof. Let us write
$$\tilde{J}_{[tn],n}^{(\mathrm{w})} - J_n =$$

$$\sum_{j=0}^{[tn]-1} (I - \mu \Sigma([tn]/n))^{[tn]-1-j} \left[\xi_{j,n}^{(\mathrm{w})} - n^{-\beta} \left(([tn]-j)^{\beta} - ([tn]-1-j)^{\beta} \right) \boldsymbol{\theta}_{t,\beta} \right],$$
where, within this proof section, we denote

$$J_n := n^{-\beta} \sum_{j=0}^{[tn]-1} (I - \mu \Sigma([tn]/n))^{[tn]-1-j} \left(([tn] - j)^{\beta} - ([tn] - 1 - j)^{\beta}) \right) \boldsymbol{\theta}_{t,\beta}.$$

Using (18), the partial sums of the terms between brackets in the next-to-last equation satisfy, for all $0 \le j \le [tn] - 1$,

$$\left| \sum_{i=j}^{[tn]-1} \left[\xi_{i,n}^{(\mathbf{w})} - n^{-\beta} \left(([tn] - i)^{\beta} - ([tn] - 1 - i)^{\beta} \right) \boldsymbol{\theta}_{t,\beta} \right] \right| = \left| \boldsymbol{\theta}(j/n) - \boldsymbol{\theta}([tn]/n) + n^{-\beta} ([tn] - j)^{\beta} \boldsymbol{\theta}_{t,\beta} \right| \le L' n^{-\beta'} ([tn] - j)^{\beta'}.$$

Integration by parts with this bound and (82), and then Lemma 33 give that there exists a constant M such that, for all $\mu \in (0, \mu_0]$ and $n \ge 1$,

$$\left| \tilde{J}_{[tn],n}^{(w)} - J_n \right| \le M(\mu \, n)^{-\beta'}.$$

Now, from (82) and using Lemma 33, we have, for all $\mu \in (0, \mu_0]$ and $n \ge 1$,

$$\left| J_n - n^{-\beta} S_{\beta} (I - \mu \Sigma([tn]/n)) \boldsymbol{\theta}_{t,\beta} \right| \le M n^{-\beta} (1 - \delta \mu)^{[tn]} \sum_{l \ge 1} (1 - \delta \mu)^l l^{\beta - 1} \le M (1 - \delta \mu)^{[tn]} (\mu n)^{-\beta},$$

where $S_{\beta}(A) := \sum_{i=0}^{\infty} A^{i} \left((i+1)^{\beta} - i^{\beta} \right)$. Using that $A \to S_{\beta}(A)$ is a power series with unit radius of convergence, (82) and (86), from the mean value theorem and Lemma 33, we have, for all $\mu \in (0, \mu_{0}]$ and all $n \geq 1$,

$$|S_{\beta}(I - \mu \Sigma([tn]/n)) - S_{\beta}(I - \mu \Sigma(t))| \le M \mu n^{-\beta} \sum_{i>1} (1 - \delta \mu)^i i^{\beta} \le M (\mu n)^{-\beta}.$$

Collecting the last three inequalities together with Lemma 33, we obtain the result. $\hfill\Box$

Proposition 25. Assume (A1) with $q \ge 4$. Let $\beta \in (0,1]$, L > 0, $\rho < 1$, and $0 < \sigma_{-} \le \sigma_{+}$. Then, there exist constants M > 0 and $\mu_{0} > 0$ such that, for all $(\theta, \sigma) \in C^{\star}(\beta, L, \rho, \sigma_{-}, \sigma_{+})$, $\mu \in (0, \mu_{0}]$ and $1 \le k \le n$,

$$\left| \mathbb{E}_{\boldsymbol{\theta},\sigma} \left[\tilde{J}_{k,n}^{(v)} \tilde{J}_{k,n}^{(v)T} \right] - \mu \frac{\sigma^2(k/n)}{2} I \right| \le M \mu \left(\mu + (\mu n)^{-\beta} + (1 - \delta \mu)^k \right)$$

Proof. Since $\{\xi^{(v)}_{i,n}, j \geq 0\}$ is a martingale increment sequence, for any $1 \leq k \leq n$, $\mathbb{E}\left[\tilde{J}_{k,n}^{(v)}\tilde{J}_{k,n}^{(v)T}\right]$ reads

$$\mu^{2} \sum_{j=0}^{k-1} (I - \mu \Sigma(k/n))^{k-1-j} \sigma_{j,n}^{2} \mathbb{E}[L_{\mu}(\mathbf{X}_{j,n}) L_{\mu}^{T}(\mathbf{X}_{j,n})] (I - \mu \Sigma(k/n))^{k-1-j}$$
$$= \mu \sigma_{k,n}^{2} (G_{k,n} - \tilde{G}_{k,n}) + R_{k,n},$$

where, for all $1 \le k \le n$,

$$G_{k,n} := \mu \sum_{j=-\infty}^{k-1} (I - \mu \Sigma(k/n))^{k-1-j} \Sigma(k/n) (I - \mu \Sigma(k/n))^{k-1-j},$$

$$\tilde{G}_{k,n} := \mu \sum_{j=-\infty}^{-1} (I - \mu \Sigma(k/n))^{k-1-j} \Sigma(k/n) (I - \mu \Sigma(k/n))^{k-1-j},$$

$$|R_{k,n}| \le M \mu^2 \sum_{j=0}^{k-1} (1 - \delta \mu)^{2(k-1-j)} (\tau^j + n^{-\beta} (k-j)^{\beta} + \mu)$$

$$\le M \mu \left(\mu + (\mu n)^{-\beta}\right).$$

For bounding $R_{k,n}$, we have used (82), (84), $\sigma \in \Lambda(\beta, L)$ and then Lemma 33. From (82) and (85), we have, for all $\mu \in (0, \mu_0]$, $1 \le k \le n$,

(94)
$$\left| \tilde{G}_{k,n} \right| \le M \left(1 - \delta \, \mu \right)^k.$$

From the previous bounds, we obtain, for all $\mu \in (0, \mu_0]$ and $1 \le k \le n$,

(95)
$$\left| \mathbb{E} \left[\tilde{J}_{k,n}^{(v)} \tilde{J}_{k,n}^{(v)T} \right] - \mu \, \sigma_{k,n}^2 \, G_{k,n} \right| \le M \, \mu \, \left(\mu + (\mu \, n)^{-\beta} + (1 - \delta \, \mu)^k \right)$$

Now, by definition of $G_{k,n}$, we have

$$(I - \mu \Sigma(k/n)) G_{k,n} (I - \mu \Sigma(k/n)) + \mu \Sigma(k/n) = G_{k,n}, \quad 1 \le k \le n,$$

which gives, for all $1 \le k \le n$,

$$(\Sigma(k/n) (I - 2G_{k,n}) + (I - 2G_{k,n}) \Sigma(k/n)) = 2 \mu \Sigma(k/n) G_{k,n} \Sigma(k/n).$$

From (94), $\sup_{1 \le k \le n} |G_{k,n}| < \infty$ uniformly over $\mu \in (0, \mu_0]$. We thus have, for all $\mu \in (0, \mu_0]$ and $1 \le k \le n$,

$$|\Sigma(k/n) (I - 2G_{k,n}) + (I - 2G_{k,n}) \Sigma(k/n)| \le M \mu.$$

For any $d \times d$ matrix C and positive definite matrix S, the equation SB + BS = C has a unique solution B linear in C and continuous in S over the set of positive definite matrices (see (Horn and Johnson, 1991, Corollary 4.4.10)). Hence we obtain, for all $\mu \in (0, \mu_0]$ and $1 \leq k \leq n$, $|I - 2G_{k,n}(\mu)| \leq M \mu$, which, with (95), gives the claimed bound.

6.4. **Proof of Theorem 6.** We use the decomposition of δ as

$$\delta^{(\mathbf{u})} + (\delta^{(\mathbf{w})} - J^{(\mathbf{w},0)})) + (\delta^{(\mathbf{v})} - J^{(\mathbf{v},0)}) + (J^{(\mathbf{w},0)} - \tilde{J}^{(\mathbf{w})}) + (J^{(\mathbf{v},0)} - \tilde{J}^{(\mathbf{v})}) + \tilde{J}^{(\mathbf{w})} + \tilde{J}^{(\mathbf{v})}.$$

Let $\eta \in (0, 1)$. Applying (53), (69), (77), (89) and (90), there exists M > 0 such that, for all $t \in [\eta, 1]$, $\mu \in (0, \mu_0]$ and $n \ge 1$,

$$\|\widehat{\boldsymbol{\theta}}_{n}(t;\mu) - \boldsymbol{\theta}(t) - \tilde{J}_{[tn],n}^{(w)} - \tilde{J}_{[tn],n}^{(v)}\|_{2,\boldsymbol{\theta},\sigma} \leq M \left(\sqrt{\mu} (\mu n)^{-\beta} + (\mu n)^{-2\beta} + \mu\right).$$

We then obtain (19) by applying Proposition 24 and Proposition 25, and using that σ is β -Lipschitz to approximate $\sigma^2([tn]/n)$ by $\sigma^2(t)$.

6.5. **Proof of Theorem 7.** We use that $\delta = \delta^{(u)} + \delta^{(v)} + (\delta^{(w)} - J^{(w,0)}) + (J^{(w,0)} - \tilde{J}^{(w)}) + \tilde{J}^{(w)}$. Observe that there exists C > 0 such that $\Lambda(\beta, L) \subseteq \Lambda(1, CL)$. Hence we may apply (53), (59), (69) and (89), so that, there exists M and μ_0 such that, for all $\mu \in (0, \mu_0]$

(96)
$$\sup_{(\boldsymbol{\theta},\sigma)\in\mathcal{C}^{\star}} \sup_{t\in[\eta,1]} \|\widehat{\boldsymbol{\theta}}_n(t;\mu) - \boldsymbol{\theta}(t) - \tilde{J}_{[tn],n}^{(w)}\|_{p,\boldsymbol{\theta},\sigma} \le M\left(\sqrt{\mu} + (\mu n)^{-2}\right).$$

Now, since $\beta > 1$, for all $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}^*$ and $t \in (0, 1]$, we may apply the Taylor expansion $\boldsymbol{\theta}(u) = \boldsymbol{\theta}(t) + \dot{\boldsymbol{\theta}}(v)(u-t)$, where $v \in [u, t]$, which yields $|\boldsymbol{\theta}(u) - \boldsymbol{\theta}(t) + \dot{\boldsymbol{\theta}}(t)(t-u)| \leq |\dot{\boldsymbol{\theta}}(v) - \dot{\boldsymbol{\theta}}(t)| |u-t| \leq L |t-u|^{\beta-1}$. Hence (18) holds $(\beta = 1 \text{ and } \beta' \text{ equal to the actual } \beta)$ at every point t > 0 and we may apply Proposition 24 for computing $\tilde{J}^{(w)}$, which easily yield the result.

APPENDIX A. PROOF OF THEOREM 16

We first derive two simple lemmas, valid under the assumptions of Theorem 16. We let ν_0 and δ denote some constants only depending on r, R_1 , ν_1 , α_1 , C_1 , λ , B and s and we write \mathbb{E}_k for $\mathbb{E}^{\mathcal{F}_k}$.

Lemma 26. For any $a \ge 1$, there exist $\delta > 0$ and $\nu_0 > 0$ such that, for all $k \in \mathbb{N}$ and $\nu \in [0, \nu_0]$,

(97)
$$\mathbf{I}(\phi_k \le R_1) \, \mathbb{E}_k \left\{ \left| \prod_{i=k+1}^{k+r} (I - \nu A_i) \right|^a \right\} \le e^{-\delta \nu}.$$

Proof. Under (C-1), we have $|I - \nu A_k| \le 1$ for all $k \in \mathbb{N}$ and $\nu \in [0, \nu_1]$ (see Lemma 31) so that we may assume a = 1 without loss of generality. We write

(98)
$$\prod_{l=k+1}^{k+r} (I - \nu A_l) = I - \nu D_k + S_k,$$

where

$$D_k := \sum_{l=k+1}^{k+r} A_l$$
 and $S_k := \sum_{j=2}^r (-1)^j \nu^j \sum_{1 \le i_1 < \dots < i_j \le r} A_{k+i_j} \dots A_{k+i_1}.$

For $\beta \in (1/s, 1)$ where s is defined in (C-3) and $\nu \geq 0$ denote

$$\mathsf{B}_k(\nu) := \{ |A_{k+1}| \le \nu^{-\beta}, \dots, |A_{k+r}| \le \nu^{-\beta} \}$$

and $\mathsf{B}_{k}^{c}(\nu)$ its complementary set. From (C-1), we have that, for all $\nu \in (0,\nu_{1}], |\nu D_{k}| \leq r\nu^{1-\beta}\mathbf{I}(\mathsf{B}_{k}(\nu)) + r\mathbf{I}(\mathsf{B}_{k}^{c}(\nu))$. Choosing $\nu_{2} \in (0,\nu_{1}]$ such that $r\nu_{2}^{1-\beta} \leq 1$, we get that, for all $\nu \in [0,\nu_{2}], 2|\nu D_{k}|\mathbf{I}(|\nu D_{k}| > 1) \leq 2r\mathbf{I}(\mathsf{B}_{k}^{c}(\nu))$. Hence, using (98) and Lemma 31, we obtain, for all $\nu \in [0,\nu_{2}]$,

$$\left| \prod_{l=k+1}^{k+r} (I - \nu A_l) \right| \le 1 - \nu \lambda_{\min}(D_k) + 2r \mathbf{I}(\mathsf{B}_k^c(\nu)) + |S_k|.$$

Eq (97) easily follows from this bound with (C-2) and the two following inequalities, which will be shown to hold for all $\nu \in [0, \nu_1]$,

(99)
$$\mathbf{I}(\phi_k \le R_1) \mathbb{E}_k[\mathbf{I}(\mathsf{B}_k^c(\nu))] \le C_1 \nu^{s\beta},$$

(100)
$$\mathbf{I}(\phi_k \le R_1) \, \mathbb{E}_k[|S_k|] \le M \, C_1^{(s \wedge 2)/s} \, \nu^{s \wedge 2},$$

where M is some constant only depending on r. We now conclude the proof by showing these two last inequalities successively.

Using the Markov inequality we obtain $\mathbb{E}_{k}[\mathbf{I}(\mathsf{B}_{k}^{c}(\nu))] \leq \sum_{l=k+1}^{k+r} \mathbb{P}_{k}\{|A_{l}| > \nu^{-\beta}\} \leq \nu^{s\beta} \sum_{l=k+1}^{k+r} \mathbb{E}_{k}[|A_{l}|^{s}]$, which implies (99) using (C-3).

For all j = 2, ..., r, for all ordered j-tuple $1 \le i_1 < \cdots < i_j \le r$, using (C-1), we have, for all $\nu \in [0, \nu_1]$, $\nu^j \left| A_{k+i_1} ... A_{k+i_j} \right| \le \nu^2 \left| A_{k+i_1} A_{k+i_2} \right|$. Hence, for some constant M_1 only depending on r, for all $\nu \in [0, \nu_1]$,

$$\mathbb{E}_k[|S_k|] \le M_1 \nu^2 \sup_{1 \le i < j \le r} \mathbb{E}_k[|A_{k+i}A_{k+j}|].$$

Put $\tilde{s} = s \wedge 2$. The Hölder inequality gives

$$\mathbb{E}_{k}[|A_{k+i}A_{k+j}|] \leq \{\mathbb{E}_{k}[|A_{k+i}|^{\tilde{s}}]\}^{\frac{1}{\tilde{s}}}\{\mathbb{E}_{k}[|A_{k+j}|^{\frac{\tilde{s}}{\tilde{s}-1}}]\}^{\frac{\tilde{s}-1}{\tilde{s}}}.$$

Observing that $\frac{\tilde{s}}{\tilde{s}-1} - \tilde{s} \geq 0$ and using (C-1), we have, for all $\nu \in [0, \nu_1]$, $\mathbb{E}_k[|A_{k+j}|^{\frac{\tilde{s}}{\tilde{s}-1}}] \leq \mathbb{E}_k[|A_{k+j}|^{\tilde{s}}]\nu^{-\frac{\tilde{s}}{\tilde{s}-1}+\tilde{s}}$, showing

$$\mathbb{E}_{k}[|S_{k}|] \leq M_{1} \nu^{\tilde{s}} \left\{ \sup_{1 \leq i < r} \mathbb{E}_{k}[|A_{k+i}|^{\tilde{s}}] \right\}^{\frac{1}{\tilde{s}}} \left\{ \sup_{1 < j \leq r} \mathbb{E}_{k}[|A_{k+j}|^{\tilde{s}}] \right\}^{\frac{\tilde{s}-1}{\tilde{s}}} \\
\leq M_{1} \nu^{\tilde{s}} \sup_{1 \leq i \leq r} \mathbb{E}_{k}[|A_{k+i}|^{\tilde{s}}] \leq M_{1} \nu^{\tilde{s}} \left(\sup_{1 \leq i \leq r} \mathbb{E}_{k}[|A_{k+i}|^{s}] \right)^{\tilde{s}/s}$$

The proof of (100) then follows by bounding the above sup by a sum and by applying (C-3). \Box

Define $N_n := \sum_{l=0}^{\lfloor n/r \rfloor} \mathbf{I}(\phi_{lr} \leq R_1)$. We have

Lemma 27. There exist $\alpha_0 > 0$ and $\gamma > 0$ such that, for all $n \geq 0$,

$$\mathbb{E}_0\left[e^{-\alpha_0 N_n}\right] \le e^{-\gamma n \alpha_0} V_0.$$

Proof. Observe that, using (C-4), for all $k \in \mathbb{N}$,

$$\mathbb{E}_{kr}\left[V_{(k+1)r}\right] \leq V_{kr}\left[\lambda \mathbf{I}(\phi_{kr} > R_1) + B\mathbf{I}(\phi_{kr} \leq R_1)\right] = V_{kr}\lambda^{\mathbf{I}(\phi_{kr} > R_1)} B^{\mathbf{I}(\phi_{kr} \leq R_1)}.$$

Let $\{W_{kr}, k \in \mathbb{N}\}$ be the process defined by

$$W_0 := V_0$$
 and $W_{kr} = \left(\frac{1}{\lambda}\right)^k \left(\frac{\lambda}{B}\right)^{N_{(k-1)r}} V_{kr}, k \ge 1.$

Since N_{kr} is \mathcal{F}_{kr} -measurable, we obtain, for all $k \in \mathbb{N}$,

$$\mathbb{E}_{kr}[W_{(k+1)r}] = \left(\frac{1}{\lambda}\right)^{k+1} \left(\frac{\lambda}{B}\right)^{N_{kr}} \mathbb{E}_{kr}[V_{(k+1)r}]$$

$$\leq \left(\frac{1}{\lambda}\right)^{k} \left(\frac{\lambda}{B}\right)^{N_{kr} - \mathbf{I}(\phi_{kr} \leq R_1)} V_{kr} = W_{kr}.$$

Hence, by induction, $\mathbb{E}_0[W_{kr}] \leq W_0 = V_0$ and, since $V_k \geq 1$, we get

$$\mathbb{E}_{0}\left[\left(\frac{\lambda}{B}\right)^{N_{kr}}\right] \leq \lambda^{k+1}\mathbb{E}_{0}\left[\left(\frac{1}{\lambda}\right)^{k+1}\left(\frac{\lambda}{B}\right)^{N_{kr}}V_{(k+1)r}\right]$$
$$\leq \lambda^{k+1}\mathbb{E}_{0}[W_{(k+1)r}] \leq \lambda^{k+1}V_{0}.$$

Noting that $N_{kr+q} = N_{kr}$ for all $q = 0, 1, \dots, r-1$, the proof follows.

We now turn back to the proof of Theorem 16. From (C-1), for all $k \in \mathbb{N}$ and $\nu \in [0, \nu_1]$,

$$(101) \left| \prod_{l=kr+1}^{(k+1)r} (I - \nu A_l) \right|^p \le e^{-(\delta/2)\nu \mathbf{I}(\phi_{kr} \le R_1)} \left\{ \left| \prod_{l=kr+1}^{(k+1)r} (I - \nu A_l) \right|^p e^{(\delta/2)\nu} \mathbf{I}(\phi_{kr} \le R_1) + \mathbf{I}(\phi_{kr} > R_1) \right\},$$

where δ is defined in Lemma 26. Let n=mr+t, where $m\in\mathbb{N}$ and $t=0,1,\ldots,r-1$. Eq. (101) and the Cauchy-Schwarz inequality show that

$$\mathbb{E}_0 \left| \prod_{l=1}^n (I - \nu A_l) \right|^p \le \pi_1^{1/2} \, \pi_2^{1/2}, \quad \text{where} \quad \pi_1 = \mathbb{E}_0 \left[e^{-\delta \nu N_n} \right] \quad \text{and} \quad$$

$$\pi_2 = \mathbb{E}_0 \left[\prod_{k=0}^{m-1} \left\{ \left| \prod_{l=kr+1}^{(k+1)r} (I - \nu A_l) \right|^{2p} e^{\delta \nu} \mathbf{I}(\phi_{kr} \le R_1) + \mathbf{I}(\phi_{kr} > R_1) \right\} \right].$$

Let $U_0 = 1$ and define recursively U_{k+1} for $k = 0, 1, \ldots$,

$$U_{k+1} := \left\{ \left| \prod_{l=kr+1}^{(k+1)r} (I - \nu A_l) \right|^{2p} e^{\delta \nu} \mathbf{I}(\phi_{kr} \le R_1) + \mathbf{I}(\phi_{kr} > R_1) \right\} U_k.$$

Applying Lemma 26 with a=2p, we obtain that (U_k, \mathcal{F}_{kr}) is a supermartingale. Consequently $\pi_2 \leq 1$. Lemma 27 and the Jensen's inequality show that, for all $\nu \in [0, \alpha_0/\delta]$,

$$\mathbb{E}_0[e^{-\delta\nu N_n}] \le \left(\mathbb{E}_0[e^{-\alpha_0 N_n}]\right)^{\delta\nu/\alpha_0} \le e^{-\gamma\delta\nu n}V_0,$$

which concludes the proof.

APPENDIX B. BURKHÖLDER INEQUALITIES FOR THE TVAR PROCESS

Throughout this section we let $\beta \in (0,1]$, L, $0 < \rho < 1$, $0 < \sigma_{-} \leq \sigma_{+}$ and we set $\mathcal{C} := \mathcal{C}(\beta, L, \rho, \sigma_{-}, \sigma_{+})$. We further let $\tau \in (\rho, 1)$. The following Lemma is adapted from (Dedecker and Doukhan, 2003, Proposition 4).

Lemma 28. Let $(\Omega, \mathcal{F}, \mathbb{P}, \{\mathcal{F}_n; n \in \mathbb{N}\})$ be a filtered space. Let $p \geq 2$, $p_1, p_2 \in [1, \infty]$ such that $p_1^{-1} + p_2^{-1} = 2p^{-1}$ and let $\{Z_n; n \in \mathbb{N}\}$ be an adapted sequence such that $\mathbb{E}[Z_n] = 0$ and $\|Z_n\|_p < \infty$ for all $n \in \mathbb{N}$. Then

(102)
$$\left\| \sum_{i=1}^{n} Z_{i} \right\|_{p} \leq \left(2p \sum_{i=1}^{n} \|Z_{i}\|_{p_{1}} \sum_{j=i}^{n} \left\| \mathbb{E}^{\mathcal{F}_{i}} \left[Z_{j} \right] \right\|_{p_{2}} \right)^{\frac{1}{2}}.$$

Proposition 29. Assume (A1) with $q \geq 2$ and let $p \geq 0$ such that $2(p+1) \leq q$. Then there exists M > 0 such that, for all $(\theta, \sigma) \in \mathcal{C}$, $1 \leq s \leq t \leq n$ and sequence $\{\phi_i\}_{s \leq i \leq t}$ in $\text{Li}(\mathbb{R}^d, 1, \mathbb{R}; p)$,

$$\left\| \sum_{i=s}^{t} \left(\phi_i(\mathbf{X}_{i,n}) - \mathbb{E}_{\boldsymbol{\theta},\sigma}[\phi_i(\mathbf{X}_{i,n})] \right) \right\|_{\frac{q}{p+1},\boldsymbol{\theta},\sigma}^2 \le M \sup_{i \in \{s,\dots,t\}} |\phi_i|_{\mathrm{Li}(p)} \sum_{i=s}^{t} |\phi_i|_{\mathrm{Li}(p)}.$$

Proof. Let us apply Lemma 28 with $Z_i = \phi_i(\mathbf{X}_{i,n}) - \mathbb{E}_{\boldsymbol{\theta},\sigma}[\phi_i(\mathbf{X}_{i,n})]$ and $p_1 = p_2 = q/(p+1)$. From the L^q stability, we see that $\|Z_i\|_{\frac{q}{p+1},\boldsymbol{\theta},\sigma} \leq M \, |\phi_i|_{\mathrm{Li}(p)}$. It now remains to bound $\left\|\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}}[Z_k]\right\|_{\frac{q}{p+1},\boldsymbol{\theta},\sigma}$. Using the exponential stability and the L^q stability, Proposition 11 shows that, for all $s \leq i \leq k \leq t$, $(\boldsymbol{\theta},\sigma) \in \mathcal{C}$, $\left\|\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}}[Z_k]\right\|_{\frac{q}{p+1},\boldsymbol{\theta},\sigma} \leq M \, \tau^{k-i} \, |\phi_k|_{\mathrm{Li}(p)}$. The proof follows. \square

Another application of Lemma 28 is the following result.

Proposition 30. Assume that $q \geq 5$ and let $p, r \geq 0$ such that $u := 2(p+r) + 5 \leq q$. There exists a constant M such that, for all $(\theta, \sigma) \in C$, $1 \leq s \leq t \leq n$ and sequences $\{\gamma_{i,j}\}_{s \leq i < j \leq t}$ and $\{\phi_i\}_{s < i \leq t}$ respectively in $\text{Li}(\mathbb{R}, 1, \mathbb{R}; p)$ and

 $Li(\mathbb{R}, 1, \mathbb{R}; r),$

$$(103) \left\| \sum_{s \leq i < j \leq t} \gamma_{i,j}(\mathbf{X}_{i,n}) \, \sigma_{i+1,n} \, \epsilon_{i+1,n} \left(\phi_{j}(\mathbf{X}_{j,n}) - \mathbb{E}_{\boldsymbol{\theta},\sigma}[\phi_{j}(\mathbf{X}_{j,n})] \right) \right\|_{\frac{2q}{u},\boldsymbol{\theta},\sigma}$$

$$\leq M \left\{ \sup_{s \leq i < j \leq t} |\gamma_{i,j}|_{\mathrm{Li}(p)} \sum_{i=s+1}^{t} |\phi_{i}|_{\mathrm{Li}(r)} + \left(\sup_{s < i \leq t} |\phi_{i}|_{\mathrm{Li}(r)} \sum_{i=s+1}^{t} |\phi_{i}|_{\mathrm{Li}(r)} \right)^{\frac{1}{2}} \sup_{s < j \leq t} \left(\sum_{i=s}^{j} |\gamma_{i,j}|_{\mathrm{Li}(p)}^{2} \right)^{\frac{1}{2}} \right\}.$$

Proof. Let $\zeta_{i,j} := \gamma_{i-1,j}(\mathbf{X}_{i-1,n}) \, \sigma_{i,n} \, \epsilon_{i,n}$ and $U_j := \phi_j(\mathbf{X}_{j,n}) - \mathbb{E}_{\boldsymbol{\theta},\sigma}[\phi_j(\mathbf{X}_{j,n})]$ for all $s < i \le j \le t$. For all $s < i \le j \le t$, U_i and $\zeta_{i,j}$ are $\mathcal{F}_{i,n}$ -measurable. All along the proof, we denote by M some constant independent of s, t, n and $(\boldsymbol{\theta}, \sigma) \in \mathcal{C}$. From uniform L^q stability, for all $s < i \le j \le t$,

(104)
$$\|\zeta_{i,j}\|_{\frac{q}{p+1},\boldsymbol{\theta},\sigma} \leq M |\gamma_{i-1,j}|_{\mathrm{Li}(p)}$$
 and $\|U_j\|_{\frac{q}{p+1},\boldsymbol{\theta},\sigma} \leq M |\phi_j|_{\mathrm{Li}(r)}$.

Denote $Z_j := U_j \sum_{i=s+1}^j \zeta_{i,j}$. The LHS of (103) then reads

(105)
$$\left\| \sum_{s < i \le j \le t} \zeta_{i,j} U_j \right\|_{\frac{2q}{u}, \boldsymbol{\theta}, \sigma}$$

$$\leq \sum_{s < i \le j \le t} \left| \mathbb{E}_{\boldsymbol{\theta}, \sigma} [\zeta_{i,j} U_j] \right| + \left\| \sum_{j=s+1}^t (Z_j - \mathbb{E}_{\boldsymbol{\theta}, \sigma} [Z_j]) \right\|_{\frac{2q}{u}, \boldsymbol{\theta}, \sigma}.$$

By Proposition 11, for all $s < i \le j < t$,

(106)
$$\left\| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[U_j \right] \right\|_{\frac{q}{\sigma-1},\boldsymbol{\theta},\sigma} \leq M \, \tau^{j-i} \, |\phi_j|_{\mathrm{Li}(r)}.$$

Using (104) and (106), for all $s < i \le j \le t$,

$$\left| \mathbb{E}_{\boldsymbol{\theta},\sigma}[\zeta_{i,j} U_j] \right| = \left| \mathbb{E}_{\boldsymbol{\theta},\sigma}[\zeta_{i,j} \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} [U_j]] \right| \leq$$

$$\left\| \zeta_{i,j} \right\|_{2,\boldsymbol{\theta},\sigma} \left\| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} [U_j] \right\|_{2,\boldsymbol{\theta},\sigma} \leq M \tau^{j-i} |\gamma_{i-1,j}|_{\mathrm{Li}(p)} |\phi_j|_{\mathrm{Li}(r)}.$$

It then follows that

$$\sum_{s < i \le j \le t} |\mathbb{E}_{\boldsymbol{\theta}, \sigma}[\zeta_{i,j} U_j]| \le M \, \gamma_{\infty, \infty}^{\dagger} \sum_{j=s+1}^{t} |\phi_j|_{\mathrm{Li}(r)} \, \sum_{i=s+1}^{j} \tau^{j-i} \le M \, \gamma_{\infty, \infty}^{\dagger} \, \phi_1^{\dagger},$$

where $\gamma_{\infty,\infty}^{\dagger} := \sup_{s \leq i < j \leq t} |\gamma_{i,j}|_{\mathrm{Li}(p)}$ and $\phi_1^{\dagger} := \sum_{i=s+1}^t |\phi_i|_{\mathrm{Li}(r)}$; Applying this bound to the RHS of (105), this yields the first term of (103).

We now bound the second term in the RHS of (105). Applying Lemma 28 with p = q/u, $p_1 = q/(p+r+2)$ and $p_2 = q/(p+r+3)$, this terms satisfies

$$\left\| \sum_{j=s+1}^{t} \overline{Z_{j}} \right\|_{\frac{2q}{\theta}, \boldsymbol{\theta}, \sigma} \leq \left(\frac{4q}{u} \sum_{j=s+1}^{t} \left\| \overline{Z_{j}} \right\|_{\frac{q}{p+r+2}, \boldsymbol{\theta}, \sigma} \sum_{k=j}^{t} \left\| \mathbb{E}_{\boldsymbol{\theta}, \sigma}^{\mathcal{F}_{j, n}} \left[\overline{Z_{k}} \right] \right\|_{\frac{q}{p+r+3}, \boldsymbol{\theta}, \sigma} \right)^{\frac{1}{2}}.$$

where $\bar{Z}_i := Z_i - \mathbb{E}_{\boldsymbol{\theta},\sigma}[Z_i]$. Applying the Hölder inequality, (104) and the Burkhölder inequality (see (Hall and Heyde, 1980, Theorem 2.12)) shows that, for all $s < j \le t$,

$$\|\overline{Z_j}\|_{\frac{q}{p+r+2},\boldsymbol{\theta},\sigma} \leq M |\phi_j|_{\mathrm{Li}(r)} \gamma_{2,\infty}^{\dagger}, \quad \text{with} \quad \gamma_{2,\infty}^{\dagger} := \sup_{s < j \leq t} \left[\sum_{i=s+1}^{j} |\gamma_{i-1,j}|_{\mathrm{Li}(p)}^2 \right]^{\frac{1}{2}}.$$

From the two last displays, we see that, in order to obtain the second term in the RHS of (103) and thus to conclude the proof, it is now sufficient to show that, for all $s < j \le t$,

$$\sum_{k=j}^{t} \left\| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j}} \left[\overline{Z_{k}} \right] \right\|_{\frac{q}{p+r+3},\boldsymbol{\theta},\sigma} \leq M \, \phi_{\infty}^{\dagger} \, \gamma_{2,\infty}^{\dagger},$$

where $\phi_{\infty}^{\dagger} := \sup_{s < i \leq t} |\phi_i|_{\mathrm{Li}(r)}$. In fact, below, we bound the LHS of the previous equation by $A_j + B_j$ and show this inequality successively for A_j and B_j . Denoting $\overline{\zeta_{i,k} U_k} := \zeta_{i,k} U_k - \mathbb{E}_{\boldsymbol{\theta},\sigma}[\zeta_{i,k} U_k]$, we have

$$\sum_{k=j}^{t} \left\| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j}} \left[\overline{Z_{k}} \right] \right\|_{\frac{q}{p+r+3},\boldsymbol{\theta},\sigma} \leq A_{j} + B_{j}, \quad \text{where}$$

$$A_{j} := \sum_{k=j}^{t} \left\| \sum_{i=s+1}^{j} \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} \left[\overline{\zeta_{i,k} U_{k}} \right] \right\|_{\frac{q}{p+r+3},\boldsymbol{\theta},\sigma},$$

$$B_{j} := \sum_{k=j}^{t} \sum_{i=j+1}^{k} \left\| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} \left[\overline{\zeta_{i,k} U_{k}} \right] \right\|_{\frac{q}{p+r+3},\boldsymbol{\theta},\sigma}.$$

The bound on A_j is obtained as follows. The centering term in the definition in A_j may be forgotten by multiplying the leading term by a factor 2. Then we use that $\zeta_{i,k} \in \mathcal{F}_{j,n}$ for all $i \leq j$ and all $k \geq i$ with the Hölder inequality, and finally apply (106) and (104) with the Burkhölder inequality

for martingale sequences to the obtained norms. These three steps read

$$A_{j} \leq 2 \sum_{k=j}^{t} \left\| \sum_{i=s+1}^{j} \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} \left[\zeta_{i,k} \, U_{k} \right] \right\|_{\frac{q}{p+r+2},\boldsymbol{\theta},\sigma}$$

$$\leq 2 \sum_{k=j}^{t} \left\| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} \left[U_{k} \right] \right\|_{\frac{q}{r+1},\boldsymbol{\theta},\sigma} \left\| \sum_{i=s+1}^{j} \zeta_{i,k} \right\|_{\frac{q}{p+1},\boldsymbol{\theta},\sigma}$$

$$\leq M \sum_{k=j}^{t} \tau^{k-j} \left| \phi_{k} \right|_{\operatorname{Li}(r)} \left(\sum_{i=s+1}^{j} \left| \gamma_{i-1,k} \right|_{\operatorname{Li}(p)}^{2} \right)^{\frac{1}{2}} \leq M \, \phi_{\infty}^{\dagger} \, \gamma_{2,\infty}^{\dagger},$$

It remains to show a similar inequality for B_j . From (104) and (106), for all $s < i \le j \le k \le t$

$$(107) \quad \left\| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[\overline{\zeta_{j,k} U_{k}} \right] \right\|_{\frac{q}{p+r+2},\boldsymbol{\theta},\sigma} \leq \left\| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} \left[\overline{\zeta_{j,k} U_{k}} \right] \right\|_{\frac{q}{p+r+2},\boldsymbol{\theta},\sigma}$$

$$\leq 2 \|\zeta_{j,k} \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} \left[U_{k} \right] \|_{\frac{q}{p+r+2},\boldsymbol{\theta},\sigma} \leq 2 \|\zeta_{j,k}\|_{\frac{q}{p+1},\boldsymbol{\theta},\sigma} \|\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{j,n}} \left[U_{k} \right] \|_{\frac{q}{r+1},\boldsymbol{\theta},\sigma}$$

$$\leq M \tau^{k-j} |\gamma_{j-1,k}|_{\mathrm{Li}(p)} |\phi_{k}|_{\mathrm{Li}(r)}.$$

This bound is in fact useful only when k-j is large. We now derive another bound of the same quantity useful when j-i is large. Since for all $i < j \le k$, $\mathbb{E}_{\boldsymbol{\theta},\sigma}[\zeta_{j,k}] = \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}}[\zeta_{j,k}] = 0$, we have

$$\mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}}\left[\overline{\zeta_{j,k}\,U_{k}}\right] = \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}}\left[\phi_{k}(\mathbf{X}_{k,n})\,\gamma_{j-1,k}(\mathbf{X}_{j-1,n})\,\sigma_{j,n}\,\epsilon_{j,n}\right] - \\ \mathbb{E}_{\boldsymbol{\theta},\sigma}\left[\phi_{k}(\mathbf{X}_{k,n})\,\gamma_{j-1,k}(\mathbf{X}_{j-1,n})\,\sigma_{j,n}\,\epsilon_{j,n}\right].$$

Note that $\sigma_{j,n}\epsilon_{j,n} = X_{j,n} - \boldsymbol{\theta}_{j-1,n}^T \mathbf{X}_{j-1,n}$ is linear in $(\mathbf{X}_{j-1,n}, \mathbf{X}_{j,n})$. By Lemma 32 and since $\sup_{j,n} |\boldsymbol{\theta}_{j-1,n}| < \infty$, $\gamma_{j-1,k}(\mathbf{X}_{j-1,n}) \phi_k(\mathbf{X}_{k,n}) \sigma_{j,n} \epsilon_{j,n}$, as a mapping of $(\mathbf{X}_{j-1,n}, \mathbf{X}_{j,n}, \mathbf{X}_{k,n})$, belongs to $\operatorname{Li}(\mathbb{R}^d, 3, \mathbb{R}; p+r+2)$ and its norm is bounded by $M |\gamma_{j-1,k}|_{\operatorname{Li}(p)} |\phi_k|_{\operatorname{Li}(r)}$. Hence, applying Proposition 11 and L^q stability to the RHS of the previous display, gives, for all $i < j \le k$,

$$\left\| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[\overline{\zeta_{j,k} U_k} \right] \right\|_{\frac{q}{n+r+3},\boldsymbol{\theta},\sigma} \leq M \, \tau^{j-i} \, |\gamma_{j-1,k}|_{\mathrm{Li}(p)} \, |\phi_k|_{\mathrm{Li}(r)}.$$

Combining with (107) we get

$$\left\| \mathbb{E}_{\boldsymbol{\theta},\sigma}^{\mathcal{F}_{i,n}} \left[\overline{\zeta_{j,k} U_k} \right] \right\|_{\frac{q}{p+r+3},\boldsymbol{\theta},\sigma} \leq M \tau^{(j-i)\vee(k-j)} |\gamma_{j-1,k}|_{\mathrm{Li}(p)} |\phi_k|_{\mathrm{Li}(r)}.$$

Applying these bounds to the definition of B_i , since, for all $i \leq t$,

$$\sum_{i < j < k < t} \tau^{(j-i)\vee(k-j)} \le 2/(\tau(1-\tau)(1-\sqrt{\tau})),$$

we finally obtain $B_i \leq M \gamma_{\infty,\infty}^{\dagger} \phi_{\infty}^{\dagger} \leq M \gamma_{2,\infty}^{\dagger} \phi_{\infty}^{\dagger}$ which achieves the proof.

APPENDIX C. TECHNICAL LEMMAS

Lemma 31. Let A be a positive semi-definite symmetric matrix and let I denote the identity matrix with same size as A. Then $|I - A| \le 1 - \lambda_{\min}(A) + 2|A|\mathbf{I}(|A| > 1)$.

Proof. Since $|\cdot|$ denotes the quadratic operator norm, we have $|I-A|=\max(1-\lambda_{\min}(A),\lambda_{\max}(A)-1)$. If $1-\lambda_{\min}(A)\geq \lambda_{\max}(A)-1$, the claimed inequality is trivially true. Since A is positive semi-definite, $\lambda_{\max}(A)=|A|$. If $1-\lambda_{\min}(A)<\lambda_{\max}(A)-1$, |I-A|=|A|-1. In addition, in this case, we necessary have |A|>1. Hence the right-hand side of the claimed inequality in this case reads $1-\lambda_{\min}(A)+2|A|=1+|A|+\lambda_{\max}(A)-\lambda_{\min}(A)\geq |A|-1$.

Lemma 32. Let $(E, |\cdot|_E)$ and $(F, |\cdot|_F)$ be two normed spaces.

(1) Let $(G, |\cdot|_G)$ be a normed space. For any $p_1, p_2 \geq 0$, there exists C > 0 such that, for all $\phi \in \text{Li}(G, 1, F; p_1)$ and $\psi \in \text{Li}(E, m, G; p_2)$,

$$|\phi \circ \psi|_{\mathrm{Li}(p_1p_2+p_1+p_2)} \le C |\phi|_{\mathrm{Li}(p_1)} (1+|\psi|_{\mathrm{Li}(p_2)}^{p_1+1}).$$

(2) Let $(\mathsf{G}, |\cdot|_{\mathsf{G}})$ be a normed algebra. For any $p_1, p_2 \geq 0$ and any integers $m_1, m_2 \geq 1$, there exists C > 0 such that, for all $\phi \in \mathrm{Li}(\mathsf{E}, m_1, \mathsf{G}; p_1)$ and $\psi \in \mathrm{Li}(\mathsf{F}, m_2, \mathsf{G}; p_2)$,

$$|\phi \psi|_{\mathrm{Li}(p_1+p_2+1)} \le C |\phi|_{\mathrm{Li}(p_1)} |\psi|_{\mathrm{Li}(p_2)}.$$

Lemma 33. Let $\beta \geq 0$ and $\nu \in (0,1)$. Then, there exists constants C_1 , C_2 only depending on β such that

$$\sup_{t>0} t^{\beta} (1-\nu)^t \le C_1 \nu^{-\beta},$$
$$\sum_{s=1}^{\infty} (1-\nu)^s s^{\beta} \le C_2 \nu^{-(1+\beta)}$$

with the convention $0^0 = 1$. Assume now that $\beta > 0$. Then, as $\nu \downarrow 0$,

(108)
$$\sum_{s=0}^{\infty} (I - \nu)^s ((s+1)^{\beta} - s^{\beta}) = \Gamma(\beta + 1) \nu^{-\beta} (1 + O(\nu)).$$

where Γ is the Gamma function.

Proof. The result is trivial for $\beta = 0$, so we assume $\beta > 0$. A straightforward computation shows that $\sup_{t \geq 0} t^{\beta} (1-\nu)^t$ is attained at $t = t_0 := -\beta/\log(1-\nu)$. Since $\log(1-\nu)$ is bounded above by $-\nu$, the first bound is obtained. The second bound is obtained by bounding $(1-\nu)^s s^{\beta}$ by this sup for $s < t_0 + 1$ and by bounding the remainder of the sum (whose terms are decreasing) by

$$\int_0^\infty (1 - \nu)^s s^\beta \, ds = \frac{1}{(-\log(1 - \nu))^{\beta + 1}} \int_0^\infty e^{-s} s^\beta \, ds,$$

which easily yields the first bound. For all $\beta > 0$ and $\nu \in (0,1)$,

$$S := \sum_{s=0}^{\infty} (1 - \nu)^s \left((s+1)^{\beta} - s^{\beta} \right) = \beta \sum_{s=0}^{\infty} (1 - \nu)^s \int_s^{s+1} t^{\beta - 1} dt.$$

The proof of (108) follows by then writing

$$(1 - \nu) S \le \beta \int_0^\infty (1 - \nu)^t t^{\beta - 1} dt = \Gamma(\beta + 1) (-\log(1 - \nu))^{-\beta} \le S.$$

References

AGUECH, R., MOULINES, E. and PRIOURET, P. (2000). On a perturbation approach for the analysis of stochastic tracking algorithm. *SIAM Journal on Control and Optimization* **39** 872–899.

Baranger, J. (1991). Analyse numérique. Hermann.

Belitser, E. (2000). Recursive estimation of a drifted autoregressive parameter. *The Annals of Statistics* **28** 860–870.

Brockwell, P. J. and Davis, R. A. (1991). *Time series: theory and methods*. 2nd ed. Springer Series in Statistics, Springer-Verlag, New York.

Dahlhaus, R. (1996a). Asymptotic statistical inference for nonstationary processes with evolutionary spectra. In *Athens Conference on Applied Probability and Time Series Analysis, Vol. II (1995)*, vol. 115 of *Lecture Notes in Statist*. Springer, New York, 145–159.

Dahlhaus, R. (1996b). On the Kullback-Leibler information divergence of locally stationary processes. *Stochastic Process. Appl.* **62** 139–168.

- Dahlhaus, R. (1997). Fitting time series models to non-stationary processes. *Annals of Statistics* **25** 1–37.
- Dahlhaus, R. and Giraitis, L. (1998). On the optimal segment length for parameter estimates for locally stationary processes. *J. Time Ser. Anal.* **19** 629–655.
- DEDECKER, J. and DOUKHAN, P. (2003). A new covariance inequality and applications. *Stochastic Process. Appl.* **106** 63–80.
- Dunford, N. and Schwartz, J. T. (1958). *Linear Operators*, vol. I. New York: : Wiley.
- GILL, R. D. and LEVIT, B. Y. (1995). Applications of the van Trees inequality: a bayesian Cramér-Rao bound. Bernoulli 59–79.
- Grenier, Y. (1983). Time-dependent ARMA modeling of nonstationary signals. *IEEE Transactions on ASSP* **31** 899–911.
- Guo, L. (1994). Stability of recursive stochastic tracking algorithms. SIAM J. on Control and Optimization 32 1195–1125.
- Hall, P. and Heyde, C. C. (1980). Martingale Limit Theory and its Application. Academic Press Inc., New York.
- Hallin, M. (1978). Mixed autoregressive-moving average multivariate processes with time-dependent coefficients. J. Multivariate Anal. 8 567–572.
- HORN, R. A. and JOHNSON, C. R. (1991). Topics in matrix analysis. Cambridge university press.
- Kailath, T. (1980). Linear Systems. Prentice Hall.
- Kushner, H. and Yin, G. (1997). Stochastic Approximation Algorithms and Applications. Springer-Verlag, New York.
- LJUNG, L. and Soderström, T. (1983). Theory and Practive of Recursive Identification. MIT Press, Cambridge, New Jersey.
- PRIOURET, P. and VERETENNIKOV, A. Y. (1995). A remark on stability of the L.M.S. tracking algorithm. *Stochastic Analysis and its Applications* **16** 118–128.
- Solo, V. and Kong, X. (1995). Adaptive Signal Processing Algorithms: Stability and Performance. Prentice Hall, New Jersey.
- Subba Rao, T. (1970). The fitting of non-stationary time-series models with time-dependent parameters. J. Roy. Statist. Soc. Ser. B 32 312–322.